

A Laboratory Medicine Best Practices Systematic Review and Meta-analysis of Nucleic Acid Amplification Tests (NAATs) and Algorithms Including NAATs for the Diagnosis of *Clostridioides (Clostridium) difficile* in Adults

 Colleen S. Kraft,^a J. Scott Parrott,^{b,c} Nancy E. Cornish,^d Matthew L. Rubinstein,^d Alice S. Weissfeld,^e Peggy McNult,^f Irving Nachamkin,^g Romney M. Humphries,^h Thomas J. Kirn,^{b,c} Jennifer Dien Bard,^{i,j} Joseph D. Lutgring,^a Jonathan C. Gullett,^k Cassiana E. Bittencourt,^l  Susan Benson,^{m,n} April M. Bobenchik,^o Robert L. Sautter,^p Vickie Baselski,^q Michel C. Atlas,^r Elizabeth M. Marlowe,^s Nancy S. Miller,^{t,u} Monika Fischer,^v Sandra S. Richter,^w Peter Gilligan,^x James W. Snyder^r

^aEmory University School of Medicine, Atlanta, Georgia, USA

^bDepartment of Interdisciplinary Studies, School of Health Professions, Rutgers University, Newark, New Jersey, USA

^cDepartment of Epidemiology, School of Public Health, Rutgers University, Piscataway, New Jersey, USA

^dCenters for Disease Control and Prevention, Atlanta, Georgia, USA

^eMicrobiology Specialists Incorporated, Houston, Texas, USA

^fAmerican Society for Microbiology, Washington, DC, USA

^gPerelman School of Medicine, University of Pennsylvania, Philadelphia, Pennsylvania, USA

^hAccelerate Diagnostics, Tucson, Arizona, USA

ⁱChildren's Hospital Los Angeles, Los Angeles, California, USA

^jKeck School of Medicine, University of Southern California, Los Angeles, California, USA

^kKaiser Permanente (Southern California Permanente Medical Group) Regional Reference Laboratories, Greater Los Angeles, Los Angeles, California, USA

^lUniversity of California—Irvine, Orange, California, USA

^mPathWest Laboratory Medicine, Perth, Western Australia, Australia

ⁿUniversity of Western Australia, Perth, Western Australia, Australia

^oRhode Island Hospital/Lifespan Academic Medical Center, Warren Alpert Medical School of Brown University, Providence, Rhode Island, USA

^pRL Sautter Consulting LLC, Lancaster, South Carolina, USA

^qUniversity of Tennessee Health Science Center, Memphis, Tennessee, USA

^rKornhauser Health Sciences Library, University of Louisville, Louisville, Kentucky, USA

^sRoche Molecular Systems, Pleasanton, California, USA

^tBoston Medical Center, Boston, Massachusetts, USA

^uBoston University School of Medicine, Boston, Massachusetts, USA

^vIndiana University, Indianapolis, Indiana, USA

^wCleveland Clinic, Cleveland, Ohio, USA

^xUniversity of North Carolina School of Medicine, Chapel Hill, North Carolina, USA

SUMMARY	2
INTRODUCTION	3
Quality Gap: Factors Associated with the Laboratory Diagnosis of <i>C. difficile</i>	3
METHODS	5
Ask: Review Questions and Analytic Framework	5
Acquire: Literature Search and Request for Unpublished Studies	6
Appraise: Screening and Evaluation of Individual Studies and Qualitative Determination of Quality and Effect	6
Quality and effect within articles	6
(i) Study risk of bias within individual articles	6
(ii) Level of effect within individual studies	7
(iii) Overall strength (level of effect) across studies.....	8
	(continued)

Citation Kraft CS, Parrott JS, Cornish NE, Rubinstein ML, Weissfeld AS, McNult P, Nachamkin I, Humphries RM, Kirn TJ, Dien Bard J, Lutgring JD, Gullett JC, Bittencourt CE, Benson S, Bobenchik AM, Sautter RL, Baselski V, Atlas MC, Marlowe EM, Miller NS, Fischer M, Richter SS, Gilligan P, Snyder JW. 2019. A laboratory medicine best practices systematic review and meta-analysis of nucleic acid amplification tests (NAATs) and algorithms including NAATs for the diagnosis of *Clostridioides (Clostridium) difficile* in adults. Clin Microbiol Rev 32:e00032-18. <https://doi.org/10.1128/CMR.00032-18>.

Copyright © 2019 Kraft et al. This is an open-access article distributed under the terms of the [Creative Commons Attribution 4.0 International license](https://creativecommons.org/licenses/by/4.0/).

Address correspondence to Colleen S. Kraft, colleen.kraft@emory.edu, or James W. Snyder, jwsnyd01@louisville.edu.

Published 29 May 2019

Analyze: Data Synthesis (Meta-analysis) and Strength of the Body of Evidence	9
Establishing LMBP practice recommendation categorization	10
Statistical analysis	10
RESULTS	10
Risk of Bias within and across Studies	11
Diagnostic Accuracy of NAAT Only and NAAT Combined with Other Tests	14
Model Diagnostics	17
Heterogeneity	17
Device as a cause of heterogeneity.....	18
Preanalytic procedure as a cause of heterogeneity: stool conforms to the container	18
Meta-analysis of NAAT Used for Repeat Testing	20
Level of Evidence across Questions	20
ADDITIONAL CONSIDERATIONS	21
Applicability and Generalizability	21
Comparison to Recent Clinical Guidelines	22
Feasibility of Implementation	22
Limitations	23
FUTURE RESEARCH	23
CONCLUSIONS	24
Practice Recommendations	24
ASM Recommendation for NAAT-Only Testing	24
ASM Recommendation for the GDH/NAAT Algorithm	24
ASM Recommendation for the GDH/toxin/NAAT Algorithm	24
ASM Recommendation for Repeated Testing Using NAAT	25
APPENDIX	25
SUPPLEMENTAL MATERIAL	26
ACKNOWLEDGMENTS	26
REFERENCES	27
AUTHOR BIOS	30

SUMMARY The evidence base for the optimal laboratory diagnosis of *Clostridioides (Clostridium) difficile* in adults is currently unresolved due to the uncertain performance characteristics and various combinations of tests. This systematic review evaluates the diagnostic accuracy of laboratory testing algorithms that include nucleic acid amplification tests (NAATs) to detect the presence of *C. difficile*. The systematic review and meta-analysis included eligible studies (those that had PICO [population, intervention, comparison, outcome] elements) that assessed the diagnostic accuracy of NAAT alone or following glutamate dehydrogenase (GDH) enzyme immunoassays (EIAs) or GDH EIAs plus *C. difficile* toxin EIAs (toxin). The diagnostic yield of NAAT for repeat testing after an initial negative result was also assessed. Two hundred thirty-eight studies met inclusion criteria. Seventy-two of these studies had sufficient data for meta-analysis. The strength of evidence ranged from high to insufficient. The uses of NAAT only, GDH-positive EIA followed by NAAT, and GDH-positive/toxin-negative EIA followed by NAAT are all recommended as American Society for Microbiology (ASM) best practices for the detection of the *C. difficile* toxin gene or organism. Meta-analysis of published evidence supports the use of testing algorithms that use NAAT alone or in combination with GDH or GDH plus toxin EIA to detect the presence of *C. difficile* in adults. There is insufficient evidence to recommend against repeat testing of the sample using NAAT after an initial negative result due to a lack of evidence of harm (i.e., financial, length of stay, or delay of treatment) as specified by the Laboratory Medicine Best Practices (LMBP) systematic review method in making such an assessment. Findings from this systematic review provide clarity to diagnostic testing strategies and highlight gaps, such as low numbers of GDH/toxin/PCR studies, in existing evidence on diagnostic performance, which can be used to guide future clinical research studies.

KEYWORDS *C. difficile* infection, diagnostic accuracy, laboratory diagnosis, meta-analysis, systematic review

INTRODUCTION

Clostridioides (Clostridium) difficile infection (CDI) is the leading cause of health care-associated infections in the United States (1, 2). It accounts for 15% to 25% of health care-associated diarrhea cases in all health care settings, with 453,000 documented cases of CDI and 29,000 deaths in the United States in 2015 (3). Acquisition of *C. difficile* as a health care-associated infection (HAI) is associated with increased morbidity and mortality. This adds a significant burden to the health care system by increasing the length of hospital stay and readmission rates, with significant financial implications. The cost of hospital-associated CDI ranges from \$10,000 to \$20,000 per case (4–7) and \$500 million to \$1.5 billion per year nationally (1, 4, 5, 8–10).

Accurate diagnosis of CDI is critical for effective patient management and implementation of infection control measures to prevent transmission (11). The diagnosis of CDI requires the combination of appropriate test ordering and accurate laboratory testing to differentiate CDI from non-CDI diarrheal cases, including non-CDI diarrhea in a *C. difficile*-colonized patient (8). Accurate diagnosis of CDI is critical for appropriate patient management and reduction of harms that may arise from diagnostic error (12) and is critical for implementation of infection control measures to prevent transmission (11). Consequently, among patients presenting with diarrhea, there is significant potential for underdiagnosis or overdiagnosis as can arise from incorrect diagnostic workups (13).

Quality Gap: Factors Associated with the Laboratory Diagnosis of *C. difficile*

Best practices for laboratory diagnosis of CDI remain controversial (14). Current laboratory practice is not standardized, with wide variation in test methods and diagnostic algorithms. Several laboratory assays are available to support CDI diagnosis in combination with clinical presentation. These include toxigenic culture (TC); the cell cytotoxicity neutralization assay (CCNA); enzyme immunoassays (EIAs) and immunochromatographic assays for the detection of glutamate dehydrogenase (GDH), toxin A or B, or both toxins; and, within the last 10 years, nucleic acid amplification tests (NAATs). Currently, two tests, TC and the CCNA, serve as reference methods for the diagnosis of *C. difficile* infection (15). The principle of the TC test is to detect strains of *C. difficile* that produce a toxin(s) following culture on an appropriate medium. CCNA detects fecal protein toxins contained within the stool and is often referred to as fecal toxin detection (16). Unfortunately, both tests are slow and labor-intensive.

Commercially available NAATs for *C. difficile* detection include those based on PCR or loop-mediated or helicase-dependent isothermal amplification (17–20). The performance of NAATs and non-NAAT tests is commonly assessed using diagnostic accuracy measures for the presence of the organism (e.g., diagnostic sensitivity, diagnostic specificity, positive predictive value [PPV], and negative predictive value [NPV]). However, these measures may not directly link to the clinical definition of CDI or clinical outcomes, and some measures (e.g., PPV and NPV) are dependent on disease prevalence in the patient population being tested (8, 17, 19, 20). Finally, in addition to diagnostic sensitivity and specificity, other factors influence the choice of testing strategy, such as cost and turnaround time.

The diagnostic accuracies of current commercially available assays (GDH EIAs, toxin A/B EIAs, and NAATs) are based on comparison with one or both of the currently accepted reference methods (TC and CCNA) for the detection of toxigenic *C. difficile*, and these comparisons are generally made to inform potential replacement of these reference methods. Although a definitive reference “gold standard” is lacking, both TC and CCNA are regarded as acceptable reference methods (15). However, some view the gold standard to be TC of a stool specimen combined with colonic histopathology of pseudomembranous colitis in patients with symptoms, but it is known that there is a spectrum of disease wherein not all patients with *C. difficile* infection have pseudomembranes (21). Finally, less frequently, colonoscopic or histopathologic findings demonstrating pseudomembranous colitis can be used in diagnostic workups to increase the diagnostic specificity for CDI diagnosis (14).

In contrasting the two reference methods (TC and CCNA), TC, while infrequently performed in clinical laboratories, is regarded as being more analytically sensitive than CCNA for detecting *C. difficile* in fecal specimens but may have lower diagnostic specificity (and, therefore, a greater likelihood of false-positive [FP] test results). CCNA has been shown to have high diagnostic sensitivity, ranging from 80 to 100%. In addition, CCNA has high diagnostic specificity and positive predictive values as well as having greater clinical utility based upon clinical outcomes (22–26). Furthermore, each reference method differs by the target detected: TC detects the presence of *C. difficile* strains that produce toxins A and/or B *in vitro* to confirm a toxigenic strain, whereas CCNA detects the presence of free toxin A or B in clinical specimens. Given these contrasting characteristics, there is potential for diagnostic discrepancy between the reference standards. Therefore, observed diagnostic performance may vary according to which reference standard is used.

Given the variety of test methods and diagnostic algorithms, there is disagreement in the laboratory community on whether best practices for the diagnosis of CDI consist of NAAT only or algorithmic testing that includes NAAT (GDH EIA followed by NAAT [GDH/NAAT] or GDH and toxin EIAs followed by NAAT [GDH/toxin/NAAT]) (20). At the initiation of these guidelines, this was the clinical quandary facing individuals who decide on a *C. difficile* testing strategy for their health care system, particularly as there is limited high-quality evidence to support which diagnostic testing strategy best supports the laboratory diagnosis of CDI (8, 22). Additionally, it remains to be determined if the potential differences in the accuracy of NAAT only or an algorithmic strategy would impact patient management or patient outcomes (27). There are few studies that encompass the nuances of laboratory CDI diagnosis as it occurs in the clinical context, for example, that evaluate the effect of preanalytic testing considerations on outcomes, to include clinical outcomes. This limitation is evident from the recent Infectious Diseases Society of America (IDSA)/Society for Healthcare Epidemiology of America (SHEA) systematic review, which included only studies that encompassed *C. difficile* testing within its clinical context, including preanalytic and postanalytic aspects (11).

Given these practice issues, and related diagnostic quality and patient safety concerns, the goal of this systematic review was to determine which laboratory testing strategies, with the inclusion of NAAT, had the best diagnostic accuracy for CDI. While it is clear that laboratory testing alone without taking into consideration the entire clinical picture is not appropriate for the diagnosis of CDI, the available literature has limited evidence linking laboratory diagnosis with clinical outcomes. Therefore, the questions for this systematic review were refined to be based only on the intermediate outcome of diagnostic accuracy for detecting the presence of the *C. difficile* organism or toxin. Although the reference standard in these studies defines what is meant by the target condition, this systematic review compares the diagnostic accuracies of these tests, including GDH detection by EIA, toxin detection by EIA, and NAAT, to those of CCNA and TC. It has been clear that preanalytical factors are crucial for NAAT specifically, and many of the studies did not include a preanalytical component, which limits whether this review can answer the question, Does this patient have *C. difficile* infection?

The questions that guided this systematic review were the following: (i) What is the diagnostic accuracy of NAAT only versus either TC or CCNA for detection of the *C. difficile* toxin gene?, (ii) What is the diagnostic accuracy of a GDH-positive EIA followed by NAAT versus either TC or CCNA for detection of the *C. difficile* organism/toxin gene?, (iii) What is the diagnostic accuracy of a GDH-positive/toxin-negative EIA followed by NAAT versus either TC or CCNA for detection of the *C. difficile* organism/toxin/toxin gene?, and (iv) What is the increased diagnostic yield of repeat testing using NAAT after an initial negative result for *C. difficile* detection of the toxin gene?

The goals of analysis based on these questions were specifically to evaluate the effectiveness of the following: (i) the diagnostic accuracies of NAAT-only and algorithmic (“two-step” or “three-step”) testing strategies, including detection of toxin or GDH

in addition to NAAT, and (ii) the diagnostic yield of repeat testing after an initial negative NAAT result. The evidence supporting these two important issues was evaluated by applying the Centers for Disease Control and Prevention (CDC) Laboratory Medicine Best Practices (LMBP) Initiative's systematic review method for translating results into evidence-based recommendations (28). The method has recently been used to evaluate practices for improving blood culture contamination (29), blood sample hemolysis (30), urine culture sample quality (31), timeliness of providing targeted therapy for bloodstream infections (32), and laboratory test utilization (33), in addition to others, and can be found at the CDC LMBP website (<https://www.cdc.gov/labbestpractices/our-findings.html>).

METHODS

This systematic review was guided by the CDC Division of Laboratory Systems (DLS) LMBP A-6 cycle, a previously validated evidence review and evaluation method for quality improvement in laboratory medicine that is reported in detail elsewhere (28). For additional resources, see www.cdc.gov/labbestpractices/index.html and <https://www.cdc.gov/library/researchguides/sytemsaticreviews.html>. The A-6 cycle was designed to assess the results of studies of practice effectiveness to derive evidence-based practice recommendations.

The systematic review was conducted by a review coordinator, a technical coordinator, a statistician with expertise in quantitative evidence analysis, and volunteer faculty (referred to as the expert panel), who were trained to apply CDC LMBP methods. The team worked with the independent, multidisciplinary LMBP Work Group consisting of 13 invited members with broad expertise in laboratory medicine, clinical practice, health services research, and health policy as well as one *ex officio* representative from the Centers for Medicare and Medicaid Services. They provided simultaneous, independent feedback on the analysis of the evidence base and the resultant American Society for Microbiology (ASM) practice recommendations. Appendix SA in the supplemental material lists the LMBP Work Group members and expert panel members and further describes their roles. The ASM Professional Practice Committee vetted the conflicts of interest for the volunteers who formed the expert panel. This systematic review and meta-analysis is intended to be updated every 5 years or as new data emerge that would change the results of this meta-analysis.

Ask: Review Questions and Analytic Framework

The overarching question addressed through this systematic review is, How effective are NAAT practices for diagnosing patients suspected of having *Clostridium difficile* infection? This review question was developed in the context of the analytic framework depicted in Fig. 1.

The "PICO" (population, intervention, comparison, outcome) elements, which directly inform study eligibility criteria, are

- Population: any patients suspected of having a *C. difficile* infection, although pediatric populations (<18 years old) were excluded from analysis
- Intervention (index test): NAAT only or diagnostic algorithms containing at least one NAAT (e.g., two-step or three-step algorithms)
- Comparison (reference method): TC and/or CCNA
- Outcomes: diagnostic accuracy (sensitivity, specificity, positive and negative likelihood ratios [LRs], and diagnostic odds ratio) of *C. difficile* testing strategies.

Preanalytic data, such as history of antibiotic use or prior hospitalization, >3 unformed bowel movements within 24 h, age of the patient, residence in long-term-care facilities, and others, are often used by health care providers when deciding if a patient should be tested for the presence of toxigenic *C. difficile*. These preanalytic data were often not included in the diagnostic accuracy comparison studies (or, if included, were not reported within the study publication), representing a study (or study reporting) limitation that should be addressed in future studies before moving forward

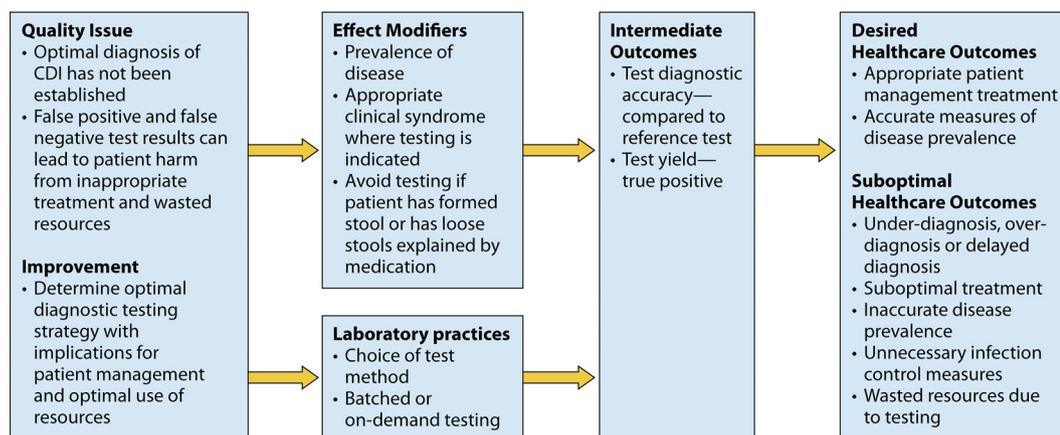


FIG 1 Analytic framework to address the question, How effective are NAAT practices for diagnosing patients suspected of having *Clostridium difficile* infection? Health care outcomes were not able to be considered in this systematic review; only intermediate outcomes could be assessed. On-demand versus batched testing could not be assessed because they were not listed in the literature.

in clinical comparison studies. See Appendix SB in the supplemental material for a suggested data collection form designed for use in future studies.

Acquire: Literature Search and Request for Unpublished Studies

The literature search strategy for eligible studies was based on input from the expert panel, a research librarian at the University of Louisville School of Medicine library in Louisville, KY, and a CDC medical librarian consultant. A search of three electronic bibliographic databases (PubMed, SCOPUS, and Embase) for English-language articles published prior to October 2016 was conducted. In addition, hand searching of bibliographies from relevant information sources was performed. Finally, solicitation of unpublished quality improvement studies was attempted by posting requests for data on both the Laboratory Medicine Best Practices website (www.cdc.gov/labbestpractices/index.html) and two ASM members-only listservs supported by the American Society for Microbiology: ClinMicroNet and ASM Division C. Further description of the search protocol as well as the full electronic search strategy for each searched database are provided in Appendix SC in the supplemental material. There was not enough literature that could be pulled during the established time frame for NAAT followed by toxin testing. This will need to be included in the update to this systematic review.

Appraise: Screening and Evaluation of Individual Studies and Qualitative Determination of Quality and Effect

Screening of search results against eligibility criteria was performed by two sets of independent reviewers, with disagreement mediated by a third reviewer.

Studies were then abstracted and quality appraised by the volunteers using a standard data abstraction form tailored to the topic of this systematic review (see Appendix SD in the supplemental material) and further adapted for use with the Systematic Review Data Repository (SRDR) online tool (<https://srdr.ahrq.gov/>). The completed data abstraction forms for each included study (referred to as “evidence summary tables”) represent consensus between two independent abstractors on content and quality appraisal, with a statistician’s review of abstracted statistical data and input of qualitative effect size ratings. Use of the data abstraction forms generated the evidence summary tables for all included studies (Appendix SE).

Studies were classified as “NAAT only,” “GDH positive (GDH⁺) EIA plus NAAT,” and “GDH⁺ and toxin negative EIA plus NAAT” (Table 1) for the purposes of this systematic review.

Quality and effect within articles. (i) Study risk of bias within individual articles. Since the primary CDC LMBP approach for evaluating study quality is not designed for

TABLE 1 Assays evaluated in this systematic review

Assay (manufacturer) ^a
NAAT only
BD GeneOhm C diff (Becton Dickinson, Sparks, MD)
Lyra Direct C diff (Quidel, San Diego, CA)
Illumigene (Meridian Bioscience, Cincinnati, OH)
Verigene (Luminex, Austin, TX)
ProGastro C. difficile (Gen-Probe Prodesse, Waukesha, WI)
Xpert C. difficile (Cepheid, Sunnyvale, CA)
Xpert C. difficile Epi (Cepheid, Sunnyvale, CA)
Portrait toxigenic C. difficile assay (Great Basin, West Valley, UT)
AdvanSure CD RT-PCR (LG Life Sciences, South Korea)
BD Max Cdiff (Becton, Dickinson, Franklin Lakes, NJ)
GDH ⁺ , NAAT
C. Diff CHEK-60 EIA (GDH) (Techlab, Blacksburg, VA) → Xpert C. difficile Epi
C. Diff CHEK-60 EIA (GDH) → Xpert C. difficile
C. Diff CHEK-60 EIA (GDH) → BD GeneOhm Cdiff assay
Quick Chek GDH (Alere, Waltham, MA) → Illumigene (Meridian Bioscience, Cincinnati, OH)
C. Diff CHEK-60 EIA (GDH) → BD GeneOhm Cdiff assay
C. Diff CHEK-60 EIA (GDH) → ProGastro CD (Prodesse, Waukesha, WI)
GDH ⁺ , toxin negative, NAAT
C. diff Quik Chek complete (Techlab, Blacksburg, VA) → GenomEra (Abacus Diagnostica, Turku, Finland)
C. diff Quik Chek complete → Xpert C. difficile
C. Diff CHEK-60 EIA (GDH) → ProSpecT C. difficile toxin A/B (Remel/Thermo Fisher, Lenexa, KS) → BD GeneOhm Cdiff assay
C. diff Quik Chek complete → Quik Chek direct (Techlab, Blacksburg, VA) → in-house PCR of <i>tcdB</i>
C. diff Quik Chek complete → Illumigene
Premier C. difficile GDH combined with ImmunoCard → Illumigene
C. diff Quik Chek complete → Prodesse ProGastro CD
C. diff Quik Chek complete → BD GeneOhm Cdiff assay

^a→ indicates a subsequent test. RT-PCR, reverse transcription-PCR.

assessing risk of bias (ROB) for diagnostic accuracy studies, the Quality Assessment of Diagnostic Accuracy Studies (QUADAS-2) tool was adapted for use with the LMBP method (34) (Table 2). Using QUADAS-2, two members of the expert panel independently assessed each study's ROB, and applicability to this review's topic, in relation to four domains: patient selection, index test, reference standard, and study flow and timing.

However, the QUADAS-2 tool provides no direct means to derive an LMBP qualitative quality rating, which is based on quality point assignment and is an essential component for deriving an LMBP practice recommendation (as described below). To meet this challenge, a key adaptation of QUADAS-2 was quality point assignment as follows: analyst responses to QUADAS-2 risk-of-bias and applicability signaling questions were categorized as 1 for yes and as 0 for either unclear or no. In the absence of the capacity for analysts to discuss risk-of-bias decisions, rating disagreements among analysts on the QUADAS-2 questions were resolved in the following manner: (i) ratings for the signaling and applicability questions were coded as 0 for "no," 0.5 for "unclear," and 1 for "yes," and (ii) ratings were averaged. The averaged risk-of-bias rating was then summed for each study. In applying the CDC LMBP method, scores of 8 to 10 received a quality rating of "good," scores of 5 to 7 received a quality rating of "fair," and scores of ≤ 4 received a quality rating of "poor" (34). In accord with the CDC LMBP method, studies receiving a poor quality rating are excluded from subsequent qualitative and quantitative syntheses. The list of arms by analysis is shown in Appendix SF in the supplemental material.

(ii) **Level of effect within individual studies.** Since diagnostic accuracy studies report two related effects (sensitivity and specificity), an approach was needed to capture the trade-off between these two measures of effect as well as the clinical

TABLE 2 Questions from QUADAS-2 used by the expert panel to evaluate studies^a

Domain	Patient selection	Index test	Reference standard	Flow and timing
Description	Describe methods of patient selection; describe included patients (prior testing, presentation, intended use of index test, and setting)	Describe the index test and how it was conducted and interpreted	Describe the reference standard and how it was conducted and interpreted	Describe any patients who did not receive the index test(s) and/or reference standard or who were excluded from the 2-by-2 table ^b ; describe the time interval and any interventions between index test(s) and reference standard
Signaling question (yes/no/unclear)	Was a consecutive or random sample of patients enrolled?	Were the index test results interpreted without knowledge of the results of the reference standard?	Is the reference standard likely to correctly classify the target condition?	Was there an appropriate interval between index test(s) and reference standard?
Risk of bias (high/low/unclear)	Was a case-control design avoided?	If a threshold was used, was it prespecified?	Were the reference standard results interpreted without knowledge of the results of the index test?	Did all patients receive a reference standard?
Concerns regarding applicability (high/low/unclear)	Did the study avoid inappropriate exclusions?	Are there concerns that the index test, its conduct, or its interpretation differed from the review question?	Are there concerns that the target condition as defined by the reference standard does not match the review question?	Did all patients receive the same reference standard?

^aAdapted from reference 34 with permission of the publisher.

^bSee the flow diagram in reference 34.

meaning of this trade-off. Additionally, an approach for deriving a single qualitative effect size rating from these measures was needed, as a necessary step when using the CDC LMBP method. The solution was based on two diagnostic accuracy effect measures: the positive likelihood ratio (+LR) (true-positive rate/false-positive rate) and the negative likelihood ratio (−LR) (false-negative rate/true-negative rate). Furthermore, the approach adopts cutoff points described by Deeks and Altman as providing a test's ability to rule in or rule out a disease and extends them into the following +LR and −LR pairings (35):

- Substantial effect rating, if +LR is >10 and −LR is <0.1
- Moderate effect rating, if +LR is >10 and −LR is >0.1 or +LR is <10 and −LR is <0.1
- Minimal effect rating, if +LR is <10 and −LR is >0.1.

In other words, the cutoffs represent thresholds for “high” clinical validity, or a “high” test information value (e.g., for determinations of posttest probability of disease for individual patients), creating a basis for judging effect sizes.

The last step in deriving a single qualitative effect size rating for each study was integrating these cutoffs into a four-quadrant likelihood ratio scatterplot of positive and negative likelihood ratio pairings, as described in the next section. In summary, this approach allows for derivation of a single effect size rating for diagnostic accuracy and allows a diagnostic accuracy evidence base synthesized using the unique qualitative synthesis approach of the CDC LMBP method.

(iii) **Overall strength (level of effect) across studies.** There are two considerations for evaluation of effect size for diagnostic accuracy statistics: (i) identifying some overall index of sample size as it relates to the interplay between sensitivity and specificity and (ii) weighing the relative risks to the patient for lower sensitivity versus lower specificity. To create an overall index of effect, the likelihood ratio scatter matrix was utilized (36).

Figure 2 demonstrates the likelihood ratio scatter matrix, which provides a practical

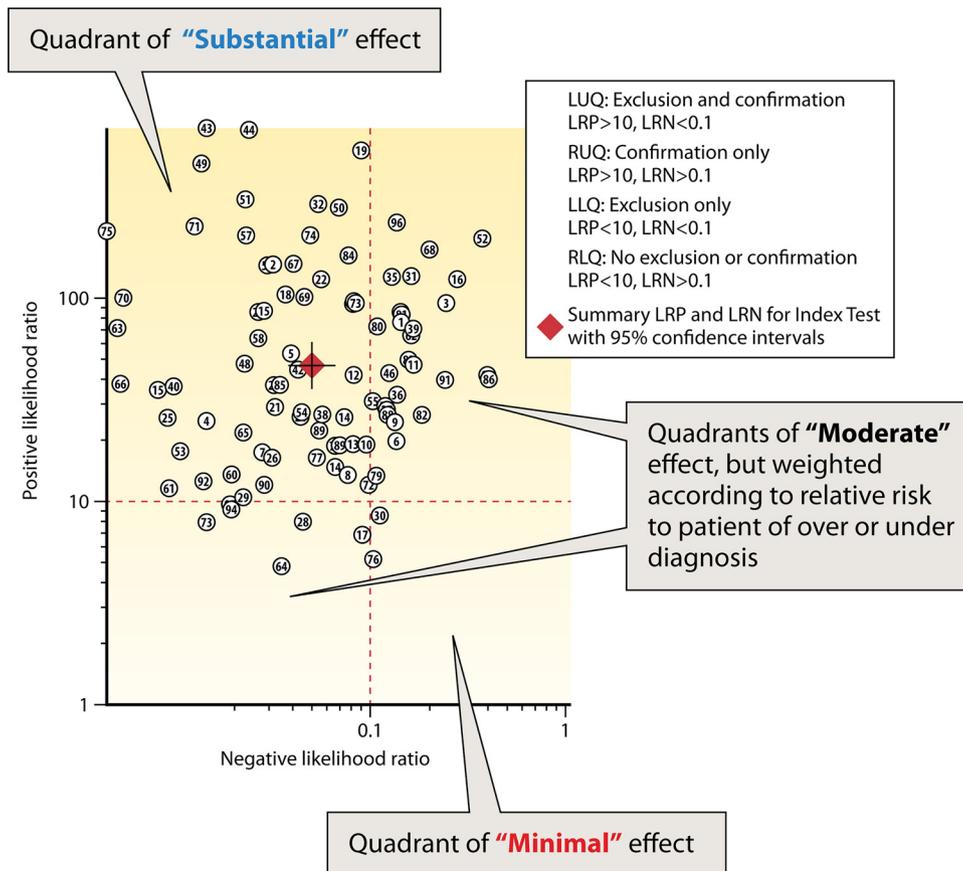


FIG 2 Example use of the likelihood ratio scatter matrix to aid in the decision of effect size. LUQ, left upper quadrant; RUQ, right upper quadrant; LLQ, left lower quadrant; RLQ, right lower quadrant; LRP, positive likelihood ratio; LRN, negative likelihood ratio.

tool to estimate the clinical validity of CDI testing approaches, based on where paired positive and negative likelihood ratios fall within the matrix quadrants (36). When paired likelihood ratios are within the areas that are typically used to indicate high clinical validity (+LR of >10 and -LR of <0.1), the expert panel could describe this as a "substantial" effect, especially if the error bands of the estimate (as represented by the cross hairs on the summary diamond) do not cross into other quadrants.

In terms of utility arising from the four combinations of -LR and +LR that are based on the likelihood ratio cutoffs described above, the scatter matrix quadrants can be further expressed as follows: the upper left quadrant signals a test that is good for both ruling in (confirming) and ruling out (excluding) a target condition, the upper right quadrant signals a test that is primarily good for ruling in, the lower left quadrant signals a test that is primarily good for ruling out, and the lower right quadrant signals a test that is not good at either ruling in or ruling out.

Analyze: Data Synthesis (Meta-analysis) and Strength of the Body of Evidence

Two analytic approaches were used in this systematic review: qualitative determinations of overall strength of evidence and quantitative meta-analysis. For the qualitative analysis, studies, grouped by testing approach, were classified according to overall strength of body of evidence, with ratings of "high," "moderate," "suggestive," or "insufficient." These qualitative ratings take into account the number of studies within a group, their effect size ratings, and their quality ratings. Criteria in Table 3 are the minimum criteria to achieve a particular LMBP strength-of-evidence rating. These criteria are the basis of the body-of-evidence qualitative analyses appearing in Results below and are the primary determinant of the best practice recommendation categorizations appearing in Conclusions below.

TABLE 3 Criteria for determining strength of body-of-evidence ratings^a

Strength of evidence	No. of studies	Effect size rating	Quality rating
High	≥3	Substantial	Good
Moderate	2	Substantial	Good
	≥3	Moderate	Good
Suggestive	1	Substantial	Good
	2	Moderate	Good
	≥3	Moderate	Fair
Insufficient	Too few	Minimal	Fair

^aAdapted from reference 28 with permission of the publisher. Also see reference 33.

For the *C. difficile* diagnostic accuracy questions, the expert panel determined that the relative harms of false-positive and false-negative results (and, therefore, of over- or underdiagnosis) are relatively equal. Although the method used to determine the body of evidence can be adjusted to fit the demands of the specific disease and diagnostic situation (36), the work group determined that this was not necessary. Thus, no adjustments were made to the basic schema pictured in Fig. 2. Finally, this approach to determining effect size applies only to diagnostic accuracy studies with outcome measures based on rates of false-positive and false-negative results (e.g., diagnostic likelihood ratios). For nondiagnostic accuracy studies (e.g., the repeat NAAT group), the review team and expert panel utilized clinical judgment to make a determination of the effect size rating. For this group, the percent diagnostic yield was determined by the percentages listed in the specific study, representing the extent to which a suspected diagnosis (CDI) was confirmed upon repeat testing by NAAT.

Establishing LMBP practice recommendation categorization. The qualitative quality ratings and qualitative effect size ratings from the individual studies for each clinical question were aggregated into bodies of evidence. The consistency of effects and patterns of effects across studies and the rating of the overall strength of the body of evidence (high, moderate, suggestive, and insufficient) were based on both qualitative and quantitative analyses using the modified LMBP process described above. Estimates of effect and the strength of the body of evidence were then used to translate results into one of three evidence-based recommendations (“recommend,” “no recommendation for or against due to insufficient evidence,” or “recommend against”). If the effect was favorable, and the overall strength of the body of evidence was either high or moderate a practice was rated as “recommended.” When the overall strength of the body of evidence was either suggestive or insufficient, a practice was rated as “no recommendation for or against due to insufficient evidence.” Categorizations of “recommend against” are used in cases where a practice is found to be antagonistic to intended outcomes (e.g., through economic outcomes, length of stay [LOS], and delay of treatment aggregated across studies in the review).

Statistical analysis. Bivariate and HSROC (hierarchical summary receiver operating characteristic) (37) models were used to estimate the summary statistics and obtain summary receiver operating characteristic (SROC) results. Analyses were carried out using the Stata 14 (Stata Corp., College Station, TX) `midas` command. Extreme outliers and highly influential cases were reevaluated and corrected by returning to the original article to determine if the values were accurate. The potential influences of quality criteria and preanalytic processes were evaluated via metaression. Model diagnostics were used to evaluate the veracity of the data. Extreme outliers and highly influential cases were reevaluated and corrected as described above if appropriate. DerSimonian-Laird models were used to estimate pooled effects of proportions of patients transitioning from negative to positive for the repeat NAAT question (38).

RESULTS

A total of 11,222 bibliographic records were identified through three electronic databases. The bibliographic records included published studies as well as conference

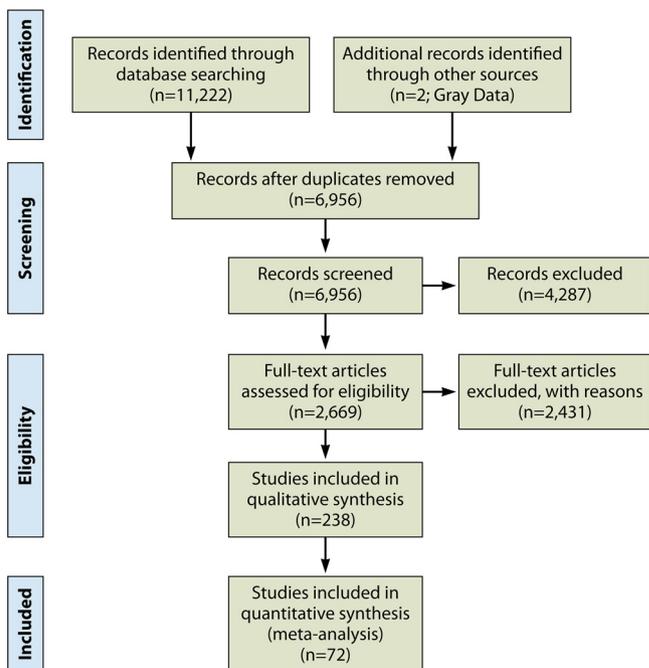


FIG 3 Study selection flow diagram.

abstracts and proceedings. Seven unpublished studies were successfully obtained for screening. After removing duplicates, a total of 6,956 bibliographic records were identified. Following the elimination of duplicate papers, the respective review and technical coordinators (J. W. Snyder and C. S. Kraft) initially screened, independently, the titles and abstracts of the 6,956 studies. Of these, 4,287 studies were excluded on the basis of not having met the following defined inclusion criteria: (i) the study did not provide valid and useful information, (ii) the patient population (>18 years of age) was not defined, (iii) there was a lack of an appropriate reference standard for comparative purposes, (iv) the study failed to address the formal study questions, (v) NAAT was not included, (vi) the article was a commentary or opinion, and (vii) the practice was not sufficiently described. Screenings of titles and abstracts independently by two reviewers (C. S. Kraft and J. W. Snyder) resulted in 2,669 studies to be considered for inclusion by full-text review. Subsequent full-text screening resulted in 238 studies determined eligible for inclusion in the systematic review, eliminating 2,431 studies not meeting inclusion criteria. Of these 238 studies, 67 could be subjected to meta-analysis due to the presence of necessary information. Five studies on the topic of repeated testing by NAAT were included, bringing the total number of studies included to 72. The study selection flow diagram is depicted in Fig. 3. Full bibliographic information for each study is provided in Appendix SE in the supplemental material.

Studies that cleared this initial screening were then abstracted by two independent reviewers and evaluated for methodological quality by the expert panel (Appendix SA). For eligible studies, information on study design, study characteristics, index and reference tests, outcome measures, and findings of the study was extracted using a standardized form adapted from the LMBP data collection tool and collected into an Agency for Healthcare Research and Quality (AHRQ)-funded online data collection platform (the Systematic Review Data Repository [<https://srd.ahrq.gov/>]) (Appendix SC).

Risk of Bias within and across Studies

Results of the risk-of-bias assessment for each of the studies included in the analyses for question 1 (NAAT only) and questions 2 and 3 (NAAT-containing algorithm) are presented in Table 4, and ROB results for question 4 (repeat testing) are reported in Table 5.

TABLE 4 QUADAS-2 risk-of-bias results and LMBP quality and effect sizes by study^a

	Risk of Bias				Applicability of Concerns			LMBP Quality Rating	LMBP Effect Size Rating
	Patient Selection	Index Test	Reference Standard	Flow and Timing	Patient Selection	Index Test	Reference Standard		
Alcala 2015 (39)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Barbady 2010 (40)	Low	Low	Unclear	Low	Low	Low	Low	Good	Substantial
Barbut 2009 (41)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Barbut 2011 (42)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Barkin 2012 (43)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Beck 2014 (44)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Brown 2011 (45)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Bruins 2012 (46)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Buchan 2012 (47)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Calderaro 2013 (48)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Carroll 2013 (49)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
de Boer 2010 (50)	Unclear	Low	Low	Low	Low	Low	Low	Good	Substantial
de Jong 2012 (51)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Devlin 2011 (52)	Low	Unclear	Low	Low	Unclear	Low	Unclear	Fair	Moderate
Doing 2012 (53)	Low	High	High	Low	Low	Low	Low	Fair	Substantial
Eastwood 2009 (54)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Ecket 2014 (55)	Low	Unclear	Low	Low	Low	Low	Low	Good	Moderate
Flore 2012 (56)	Unclear	Low	Low	Low	Low	Low	Low	Good	Substantial
Goldenberg 2010 (57)	Low	Low	Unclear	Low	Low	Low	Low	Good	Substantial
Gyorke 2013 (58)	Low	Low	High	Low	Low	Low	Unclear	Fair	Moderate
Kim 2012 (59)	Low	High	Unclear	High	Unclear	Low	Low	Fair	Substantial
Hirvonen 2013 (60)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Hoegh 2012 (61)	Low	Low	Low	Unclear	Low	Low	Low	Good	Substantial
Hong 2014 (62)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Huang 2009 (63)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Humphries 2013 (64)	Unclear	Low	Low	High	Unclear	Low	Low	Fair	Moderate
Jazmati 2015 (65)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Jensen 2015 (66)	Unclear	Low	Low	Low	Low	Low	Low	Good	Substantial
Johnstone 2010 (67)	Unclear	Unclear	Unclear	High	Unclear	Unclear	Unclear	Poor	Substantial
Karre 2011 (68)	Unclear	Low	Unclear	Low	Unclear	Low	Low	Fair	Moderate
Kato 2005 (69)	Unclear	Low	High	Unclear	High	Low	Low	Fair	Substantial
Kato 1993 (70)	High	Low	Low	Low	Unclear	Low	Low	Fair	Substantial
Kim 2012 (59)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Knetsch 2011 (71)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Kvach 2010 (72)	High	Low	High	Low	Low	Low	Low	Fair	Substantial
Kyrvenko 2010 (73)	High	Low	Low	Low	Low	Low	Low	Good	Moderate
Lalande 2011 (74)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Landry 2014 (75)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Langley 1995 (76)	Unclear	Low	Unclear	Low	Low	Low	Low	Good	Substantial
Larson 2010 (77)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Le Guern 2012 (78)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Leitner 2013 (79)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
McElgunn 2014 (80)	Low	Unclear	Low	Unclear	Unclear	Unclear	Low	Fair	Substantial
Miller 2013 (81)	Low	Low	High	High	Low	Low	Low	Fair	Substantial
Noren 2011 (82)	Unclear	Low	Low	Low	Low	Low	Low	Good	Moderate
Noren 2014 (83)	Unclear	Low	Low	Low	Low	Low	Low	Good	Substantial
Novak-Weekley 2010 (84)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Pallis 2013 (85)	Low	Low	Unclear	Unclear	Low	Low	Low	Good	Substantial

(Continued on next page)

TABLE 4 (Continued)

	Risk of Bias				Applicability of Concerns			LMBP Quality Rating	LMBP Effect Size Rating
	Patient Selection	Index Test	Reference Standard	Flow and Timing	Patient Selection	Index Test	Reference Standard		
Peterson 2011 (86)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Putsathit 2015 (87)	High	Low	High	Low	Low	Low	Low	Fair	Substantial
Shin 2012 (88)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Silva 2014 (89)	Low	Unclear	High	High	Low	High	High	Poor	Moderate
Soh 2014 (90)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Stamper 2009 (91)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Swindells 2010 (92)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
Terhes 2009 (93)	Low	Low	Unclear	Low	Low	Low	Low	Good	Substantial
Tojo 2014 (94)	Low	Low	Low	Unclear	Low	Low	Low	Good	Substantial
Van Broeck 2010 (95)	Unclear	Low	Low	Low	Unclear	Low	Low	Good	Substantial
Van Broeck 2012 (96)	Unclear	Low	Low	Low	Low	Low	Low	Good	Moderate
Vasoo 2014 (97)	Unclear	Low	Unclear	Low	Low	Low	Low	Good	Substantial
van den Berg 2005 (98)	Unclear	Low	High	Low	Low	Low	Low	Fair	Moderate
van den Berg 2006 (99)	Low	Low	Low	Low	Low	Low	Low	Good	Substantial
van den Berg 2007 (100)	Unclear	Low	Unclear	Low	Unclear	Unclear	Unclear	Fair	Substantial
Viala 2012 (101)	Unclear	Low	Low	Low	Low	Low	Low	Good	Moderate
Walkty 2013 (102)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate
Yisiurua 2013 (103)	Unclear	Low	Low	Low	Low	Low	Low	Good	Substantial
Zidaric 2011 (104)	Low	Low	Low	Low	Low	Low	Low	Good	Moderate

^aSee references 39–104.

Across studies and QUADAS-2 ROB domains included in the analysis for the first three questions, the large majority of ROB criteria were met (84.0%; 394/469). Thirty-two (47.8%) of the 67 studies had at least one ROB criterion rated as something other than a “low” risk of bias.

The two QUADAS-2 criteria in which risks were predominantly clustered were “patient selection” (31.3% of studies with some risk of bias) and “reference standard” (26.9% of studies with some risk of bias). For the patient selection criterion, the primary concern was that preanalytic procedures were not well specified. In other words, it was not made clear by the authors whether stool samples met any particular criteria (e.g., “conforms to the shape of the container”) before being included in the study, nor was it made clear whether the stool consistency was medication related, such as from laxative use. Over one-quarter of the studies (25.4%) were rated unclear for risk of bias on this criterion. Similarly, 6.0% of studies provided enough evidence to determine that the preanalytic procedures were not met (and, thus, posed a high risk of bias) for patient selection. For the repeat testing question (Table 5), patterns for risk of bias were

TABLE 5 QUADAS-2 risk-of-bias results by study for studies examined for repeat testing by NAAT^a

	Risk of Bias				Applicability of Concerns			LMBP Quality Rating	LMBP Effect Size Rating
	Patient Selection	Index Test	Reference Standard	Flow and Timing	Patient Selection	Index Test	Reference Standard		
Green 2014 (105)	High	Low	Low	Low	Low	Low	Low	Good	Minimal
Khanna 2012 (106)	Unclear	Unclear	NA	Low	Low	Low	NA	Good	Minimal
Luo 2010 (107)	Low	Low	Low	Low	Low	Low	Low	Good	Minimal
Nistico 2013 (108)	Low	Low	Unclear	Unclear	Low	Low	Unclear	Good	Minimal
Deshpande 2012 (109)	Low	Low	NA	Low	Low	Low	NA	Good	Minimal

^aSee references 105–109. NA, not applicable.

TABLE 6 Diagnostic accuracy statistics by number of tests

Parameter ^a	Value for test					
	NAAT only		GDH/NAAT		GDH/toxin/NAAT	
	Estimate	95% CI	Estimate	95% CI	Estimate	95% CI
No. of studies	96		12		9	
Prevalence	0.17		0.11		0.13	
Sensitivity	0.95	0.94–0.96	0.91	0.86–0.95	0.89	0.84–0.92
ICC SEN ^b	0.27	0.18–0.35	0.10	0.00–0.23	0.03	0.00–0.15
Specificity	0.98	0.97–0.98	0.99	0.98–1.0	0.99	0.98–1.00
ICC SPE ^c	0.27	0.19–0.34	0.25	0.00–0.53	0.26	0.00–0.62
Positive likelihood ratio	46.0	35.7–59.2	113.5	49.9–258.1	155.8	57.7–420.2
Negative likelihood ratio	0.05	0.04–0.06	0.09	0.06–0.14	0.11	0.08–0.16
Diagnostic odds ratio	934	652–1,338	1,282	484–3,395	1,383	436–4,388

^aICC, interclass correlation coefficient; SEN, sensitivity; SPE, specificity.

^bProportion of total variance in sensitivity explained by between-study variation.

^cProportion of total variance in specificity explained by between-study variation.

similar (80.6% [25/31] of criteria across studies and domains with low ROB), showing a similar weakness in the patient selection criterion.

Diagnostic Accuracy of NAAT Only and NAAT Combined with Other Tests

A total of 117 comparisons of NAAT only or NAAT-containing algorithms (GDH/NAAT or GDH/toxin/NAAT) to either TC or CCNA were extracted from 67 unique studies (Table 6). Across test arms (i.e., NAAT only, GDH/NAAT, or GDH/toxin/NAAT), the pretest probabilities (or prevalences) of the presence of *C. difficile* ranged from 11% to 17%, based on percent positive results with the reference standard, and therefore, predictive values should be interpreted accordingly. Of note, the number of comparisons within each arm differed dramatically, which substantially affects not only the accuracy of the estimates but also the confidence in the estimates. These results demonstrate that there is confidence in the diagnostic accuracy findings for the NAAT-only arm but less confidence in the exact estimates of the GDH/NAAT and GDH/toxin/NAAT arms, which is due to sample size.

There were three reference methods included in the analyses: TC, CCNA, and combined TC and CCNA. The breakdown of reference methods used by test algorithm is presented in Table 7, with counts representing individual studies.

Due to the small number of studies in the GDH/NAAT and GDH/toxin/NAAT scenarios for the different reference methods, diagnostic accuracy subgroups based on these observed differences in reference methods could not be constructed for each test algorithm approach. However, Table 8 provides diagnostic accuracy statistics grouped for each of the three reference method approaches (toxigenic culture, CCNA, and both used in combination) observed in the evidence base (Table 7). This sensitivity analysis has the tests aggregated as one comparator to each reference method. The purpose of this sensitivity analysis was to determine if the diagnostic accuracies of these assays were different if they were compared to a different reference standard.

While the specificity observed in Table 6 remained very high across arms (0.98 to 0.99), the sensitivity for detection of *C. difficile* decreased as additional tests were added prior to the NAAT, decreasing from 0.95 for NAAT only to 0.89 for GDH/toxin/NAAT.

TABLE 7 Reference method frequencies by NAAT only, GDH/NAAT, or GDH/toxin/NAAT

Reference method ^a	No. of studies			
	NAAT only	GDH/NAAT	GDH/toxin/NAAT	Total
TC	61	10	4	75
CCNA	26	2	5	33
Combined	10	0	0	10
Total	97	12	9	118

^aTC, toxigenic culture; CCNA, cell cytotoxicity neutralization assay.

TABLE 8 Sensitivity analysis of diagnostic accuracy statistics by reference standard^a

Parameter	Value					
	Toxigenic culture		CCNA		Combined TC/CCNA	
	Estimate	95% CI	Estimate	95% CI	Estimate	95% CI
No. of studies	74		33		10	
Prevalence	0.16		0.16		0.21	
Sensitivity	0.94	0.92, 0.95	0.93	0.93, 0.95	0.99	0.96, 1.00
ICC SEN ^b	0.22	0.13, 0.31	0.17	0.06, 0.28	0.39	0.03, 0.74
Specificity	0.99	0.98, 0.99	0.98	0.96, 0.98	0.98	0.96, 0.99
ICC SPE ^c	0.26	0.18, 0.35	0.30	0.17, 0.43	0.32	0.04, 0.60
Positive likelihood ratio	65.3	48.7, 87.8	38.5	24.9, 59.5	57.5	24.3, 135.9
Negative likelihood ratio	0.06	0.05, 0.08	0.08	0.05, 0.11	0.01	0.00, 0.04
Diagnostic odds ratio	1,079	745, 1,563	509	302, 857	5,022	1,127, 22,377

^aCCNA, cell cytotoxicity neutralization assay; TC, toxigenic culture; ICC, interclass correlation coefficient.

^bProportion of total variance in sensitivity explained by between-study variation.

^cProportion of total variance in specificity explained by between-study variation.

While all arms can be expected to be highly specific, there may be decreases in sensitivity when a GDH/toxin/NAAT algorithm is used. In general, a progressive decrease in overall diagnostic sensitivity (or overall sensitivities no greater than the lowest sensitivity of an individual component test) may be observed as one applies additional testing when the initial test results are positive within a combined testing algorithm. There is a progressive increase in overall diagnostic specificity (or overall specificities at least as high as the highest specificity of an individual component test) that may be observed as one progresses through such an algorithm (110).

The positive likelihood ratio (+LR) indicates how much more likely a person with the *C. difficile* organism, toxin gene, or toxin in the stool is to have a positive result on the NAAT only or NAAT algorithm than a person without the organism, toxin gene, or toxin in the stool. Typically, these ratios refer to individuals having the disease or not having the disease, but since these studies are comprised of a positive or negative test rather than predicting the disease, the LR in this review refers to presence of the organism, toxin gene, or toxin. Thus, a +LR of 46 indicates that a person with the *C. difficile* organism or toxin detected is 46 times more likely to have a positive result (Table 6) than a person who does not have the *C. difficile* organism or toxin. On the other hand, the negative likelihood ratio (−LR) indicates how less likely a person with the *C. difficile* organism or toxin detected is to have a negative result than a person without the organism or toxin. Thus, a −LR of 0.05 indicates that a person with the disease is 20 times less likely (1/0.05) to test negative on NAAT only (Table 6) than a person without the organism. A suggested rule of thumb for “high” information value of a diagnostic test (and, therefore, high clinical validity) is to have a +LR of ≥ 10 and a −LR of ≤ 0.1 (i.e., with a +LR of ≥ 10 , there is a high likelihood that the disease is present when the test result is positive, while with a −LR of ≤ 0.1 , there is a high likelihood that the disease is absent when the test result is negative) (35).

When comparing likelihood ratios across arms, all three arms have +LR in the “high” test information value range, all > 10 . We caution the reader against interpreting the differences across arms in +LRs as “more is better.” For the GDH/NAAT and GDH/toxin/NAAT arms (Table 6), the confidence intervals on the LRs are very broad (due, in part, to the smaller number of studies), and so there is less confidence in the point estimates of the +LRs for the GDH/NAAT and GDH/toxin/NAAT arms. It is justified to conclude that a positive *C. difficile* result using NAAT only or NAAT algorithms is substantially more likely in patients with the presence of the organism than in patients without the presence of *C. difficile*.

For −LR, the findings are less consistent. While the −LR for the NAAT only is 0.05, the LR for the GDH/NAAT is 0.09 (95% confidence interval [CI], 0.06 to 0.14), with the 95% confidence interval including the 0.1 cutoff.

This indicates that we cannot be entirely confident that the −LR point estimate for the GDH/NAAT arm actually meets the criterion for high information value (i.e., a high

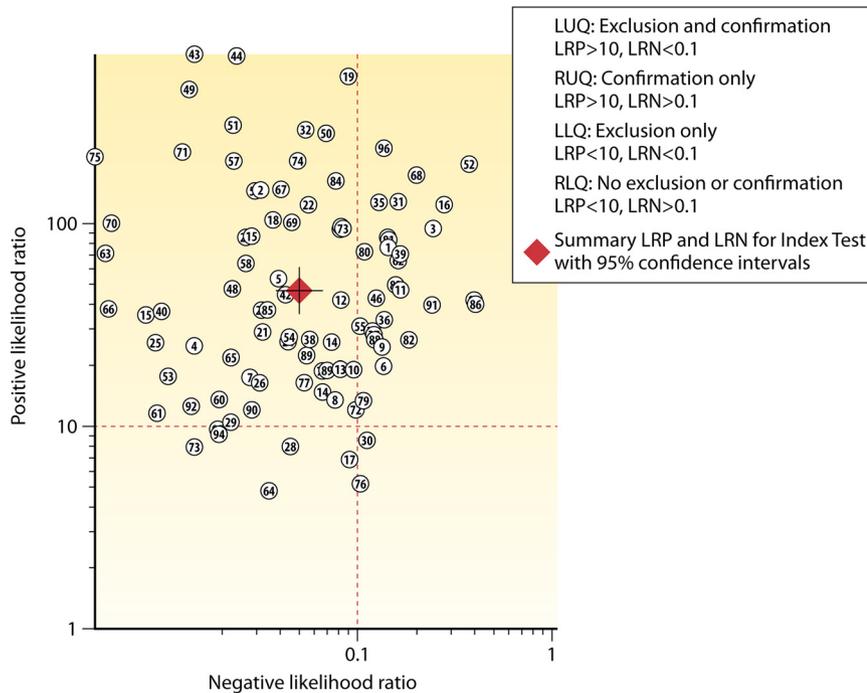


FIG 4 Scatter matrix of positive and negative likelihood ratios for NAAT-only detection of *C. difficile*. The red solid dots, in the scatter matrices, indicate the position of the combined +LR and -LR estimates. The whiskers running through the red dot are the confidence intervals for either +LR (vertical whiskers) or -LR (horizontal whiskers).

likelihood that the disease is absent when the test result is negative) (35). For the GDH/toxin/NAAT arm, the point estimate of the -LR falls above the high information value cutoff (-LR = 0.11 [95% CI, 0.08 to 0.16]), and the confidence interval indicates that the GDH/toxin/NAAT arm may be anywhere from 6.25 to 12.5 times less likely to return a negative result for a person with the presence of *C. difficile* than for a patient without the presence of *C. difficile*. In short, the false-negative rate for the GDH/toxin/NAAT arm appears to be higher than that for the NAAT-only arm and may even surpass the <0.1 high information value threshold. The reader is cautioned, however, against simply interpreting the GDH/toxin/NAAT algorithms as being less accurate than the NAAT only or GDH/NAAT, due to the small number of comparisons available for analysis. The most conservative interpretation would be that while we have confidence in the utility of the -LR to identify negative results in the NAAT-only arm, we have less confidence in the GDH/NAAT or GDH/toxin/NAAT arms, because of the small sample sizes.

The relationships between the +LR and -LR for the three arms are pictured graphically in scatter matrices (Fig. 4 to 6). The upper left quadrant of the matrices indicates the area where both +LR and -LR meet their clinical thresholds (thus, the test is useful for both excluding [i.e., accurate for true-negative results] and confirming [i.e., accurate for true-positive results] a target condition). Note that in Fig. 4 (NAAT only), not only is the summary point fully within the upper left quadrant, but the very tight confidence intervals remain within that quadrant. In contrast, while the GDH/NAAT summary point falls within the upper left quadrant (Fig. 5), the -LR confidence interval crosses into the upper right quadrant. This gives us less confidence that a GDH/NAAT diagnostic test is likely to accurately identify patients without the *C. difficile* organism or toxin gene. For the GDH/toxin/NAAT arm (Fig. 6), the summary point falls within the upper right quadrant, again causing us to be less confident that the GDH/toxin/NAAT solution can accurately identify patients without the *C. difficile* organism, toxin, or toxin gene. Study 4 in Fig. 6 is an outlier, but if excluded (data not shown), the findings would not change.

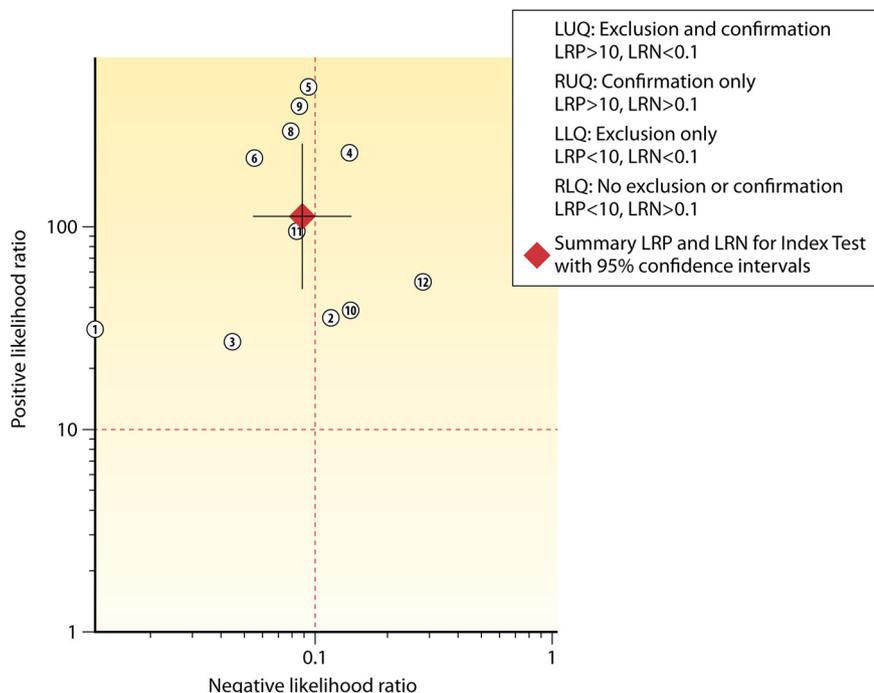


FIG 5 Scatter matrix of positive and negative likelihood ratios for GDH/NAAT algorithm detection of *C. difficile*. The red solid dots, in the scatter matrices, indicate the position of the combined +LR and -LR estimates. The whiskers running through the red dots are the confidence intervals for either +LR (vertical whiskers) or -LR (horizontal whiskers).

Model Diagnostics

Since meta-analytic procedures model the summary, estimations are only as trustworthy as the models. Thus, evaluations of model diagnostics (e.g., goodness of fit) are important (111). While there were influential outliers in both the NAAT-only and GDH/NAAT arms, these outliers were likely to pull the estimates in Table 6 slightly lower. Hence, the estimates for both the NAAT-only and the GDH/NAAT arms presented in Table 6 may be slightly conservative. However, in the GDH/toxin/NAAT arm, there was a single outlier (one study) that biased the model estimates upward, likely resulting in an overestimate of the diagnostic accuracy statistics presented in Table 6. We attempted to carry out a sensitivity analysis on the results of the GDH/toxin/NAAT arm by removing the influential case and then recomputing the statistics. However, because of the small number of comparisons, the revised model failed to achieve convergence, and no solution could be computed. Thus, readers are cautioned that while the NAAT-only and the GDH/NAAT arm estimates may be somewhat conservative (i.e., the sensitivity and specificity may be slightly higher than reported), the values for the GDH/toxin/NAAT arm may overestimate the diagnostic accuracy in this arm. Readers are encouraged to perform and contribute additional study data to add to the data available in order to refine future analyses (see Appendix SB in the supplemental material for study components needed).

Heterogeneity

Since there are variations in sample characteristics and preanalytic and analytic procedures, heterogeneity can be assumed in diagnostic accuracy meta-analyses. However, unlike meta-analyses of treatment studies, there are no generally accepted measures of heterogeneity for diagnostic accuracy meta-analyses (112). The interclass correlation coefficients (ICCs) reported in Table 6 indicate that approximately one-quarter of the differences in study findings are due to between-study differences. The analyses of heterogeneity described below are intended to explain some of the differences between studies. Two possible sources of variation were examined to

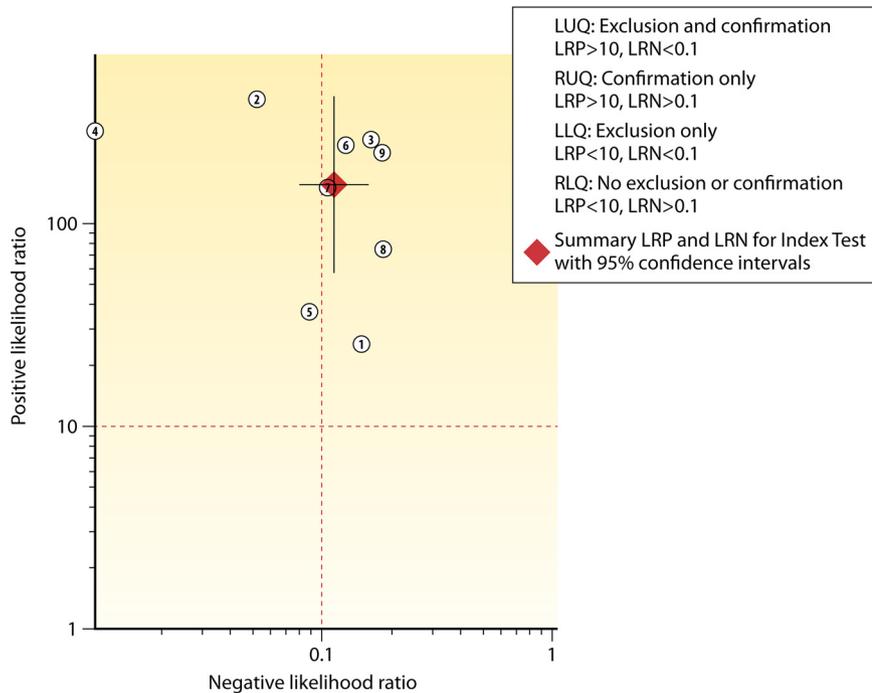


FIG 6 Scatter matrix of positive and negative likelihood ratios for GDH/toxin/NAAT algorithm detection of *C. difficile*. The red solid dots, in the scatter matrices, indicate the positions of the combined +LR and -LR estimates. The whiskers running through the red dots are the confidence intervals for either +LR (vertical whiskers) or -LR (horizontal whiskers).

determine the source of the heterogeneity and examine why some studies find higher levels of diagnostic accuracy than others. These two possible sources of variation were (i) the NAAT used and (ii) the preanalytic procedure of ensuring that the sample was unformed (i.e., conformed to the container).

Device as a cause of heterogeneity. Due to the limited number of studies available in the GDH/NAAT and GDH/toxin/NAAT arms, and because these comparisons would also introduce variation from the companion devices (e.g., GDH and toxin), we descriptively compared device (platform) differences only within the NAAT-only arm (Table 9). While there are some differences among devices, all +LR point estimates were >10, and all -LR point estimates were <0.1 (35). Only the 95% confidence interval for the GeneOhm device contained the clinical cutoff threshold for the -LR. Due to variations in other study characteristics (e.g., differences in samples or differences in analytic or preanalytic procedures), we are hesitant to conclude that one device performs better or worse than other devices. Rather, all devices appear to perform roughly the same. Thus, although variations in NAAT devices may have contributed to some variation in findings across studies, the statistics presented in Table 9 demonstrate that the variation is likely to be relatively minor.

Preanalytic procedure as a cause of heterogeneity: stool conforms to the container. The studies were assessed for whether the consistency of stool was considered in their analysis and study design. Typically, when the preanalytical aspects were involved with sample selection, only those samples that were considered to conform to the shape of the container (e.g., diarrheal, etc.) were used in the study. Analysts were directed to extract data from each of the articles on whether the authors required that specimens “conformed to the shape of the container.” Analysts could answer “yes” (authors provide a statement confirming that this preanalytic condition was met), “no” (authors give some indication that the preanalytic condition was not met), or “uncertain” (authors do not provide any information on whether the stool conformed to the container). These categories were recoded into “yes” and “no” (combining “no” and “uncertain”). For the NAAT-only and the GDH/toxin/NAAT arms, only 50% of the

TABLE 9 NAAT-only device comparison

Device and parameter	Value	
	Estimate	95% CI
GeneOhm (n = 20)		
Sensitivity	0.92	0.88–0.94
Specificity	0.98	0.97–0.99
Positive likelihood ratio	48.4	30.0–78.0
Negative likelihood ratio	0.09	0.06–0.12
Diagnostic odds ratio	569	325–996
Illumigene (n = 18)		
Sensitivity	0.95	0.93–0.97
Specificity	0.99	0.98–1.00
Positive likelihood ratio	89	40.1–197.7
Negative likelihood ratio	0.05	0.03–0.07
Diagnostic odds ratio	1,909	755–4,822
In-house (n = 12)		
Sensitivity	0.96	0.92–0.98
Specificity	0.96	0.94–0.98
Positive likelihood ratio	26.4	15.3–45.4
Negative likelihood ratio	0.04	0.02–0.09
Diagnostic odds ratio	616	233–1,630
Xpert (n = 14)		
Sensitivity	0.99	0.95–1.00
Specificity	0.97	0.94–0.98
Positive likelihood ratio	30.6	17.8–52.5
Negative likelihood ratio	0.01	0.00–0.05
Diagnostic odds ratio	3,400	611–18,920
Other device (n = 33)		
Sensitivity	0.95	0.91–0.97
Specificity	0.98	0.97–0.99
Positive likelihood ratio	47.6	30.9–73.1
Negative likelihood ratio	0.06	0.03–0.09
Diagnostic odds ratio	854	442–1,652

comparisons confirmed that the stool conformed to the container, and only one-third of the GDH/toxin cases did so (Table 10).

In studies where the authors explicitly indicated that the stool conformed to the shape of the container and tested only those that did or, its equivalent, “liquid stools” and diarrhea, etc., the diagnostic sensitivity was significantly lower than for studies where the authors did not report that the stool conformed to the shape of the container. The results were mixed for diagnostic specificity. While specificity was significantly lower in studies that reported that the stool conformed to the shape of the

TABLE 10 Comparison of sensitivities and specificities by whether authors reported that the stool conforms to the container^a

Categorization of whether stool meets criteria reported	No. of studies in arm	Sensitivity		P value for sensitivity	Specificity		P value for specificity
		Estimate	95% CI		Estimate	95% CI	
NAAT only							
Yes	48	0.94	0.92–0.96	<0.001	0.97	0.96–0.98	<0.001
No	49	0.96	0.94–0.97		0.99	0.98–0.99	
GDH/NAAT							
Yes	7	0.91	0.86–0.96	0.02	0.99	0.98–1.00	0.16
No	5	0.92	0.86–0.98		0.99	0.98–1.00	
GDH/toxin/NAAT							
Yes	4	0.86	0.79–0.92	<0.001	1.00	0.99–1.00	0.58
No	5	0.89	0.85–0.93		0.99	0.98–1.00	

^aIn those studies where the stool had to meet the criteria before being tested, only the samples that met the preanalytic requirement were tested.

TABLE 11 LMBP strength of body of evidence for all questions

Question	No. of studies	No. of comparisons	Effect	Quality
NAAT only, high strength of body of evidence	60	96	Substantial	Good
GDH/NAAT, high strength of body of evidence	9	12	Substantial	Good
GDH/toxin/NAAT, moderate strength of body of evidence	7	9	Moderate	Good
Repeat testing using NAAT, insufficient strength of body of evidence	5	6	Minimal	Good

container in the NAAT-only arm ($P < 0.001$), there was no significant difference ($P = 0.58$) for the GDH/toxin/NAAT arm, undoubtedly due to the small number of cases in this arm. In the GDH/NAAT arm, the difference showed a trend toward significance ($P = 0.16$).

In summary, the results in Table 10 indicate that the single preanalytic practice of ensuring that the stool specimen conformed to the shape of the container contributed statistically significantly to differences in sensitivity and specificity across studies. In most cases, these diagnostic accuracy estimates are higher when authors do not confirm the presence of this preanalytic practice, which is consistent with the idea that NAAT can detect *C. difficile* colonization in formed stool. Essentially, the sensitivity appears higher due to the fact that the test is detecting more positive samples (which would be considered colonization) in addition to the presence of toxigenic *C. difficile* in a diarrheal stool sample.

Meta-analysis of NAAT Used for Repeat Testing

Five studies that examined the increased diagnostic yield of NAAT used for repeat testing were available (Table 5). In this group, percent diagnostic yield represents the extent to which a suspected diagnosis (CDI) was made upon repeat testing using NAAT (an initial negative test followed by a positive test). The studies varied primarily in the length of time between repeat testing, with four studies reporting diagnostic yield over a 7-day period (105–107, 109), two studies reporting over a 14-day period (106, 107), and one study reporting over a 59-day period (108). All studies were retrospective in design. The number of repeat tests on negative samples was not reported in all studies, but for those that did report (106–108), tests were repeated between one and five times.

For the 7-day window for repeat testing by NAAT, the pooled proportion of subjects transitioning from negative to positive *C. difficile* results was 2% (95% CI, 0.009 to 0.032; $P < 0.001$). Heterogeneity was high ($I^2 = 85.23\%$; $P < 0.001$). Individual study results varied from a low of 1% transitioning (105) to a high of 3.3% (109).

For the >7-day window (14 days to 59 days), the proportion of subjects transitioning from negative to positive *C. difficile* results was 3% (95% CI, 0.023 to 0.038; $P < 0.001$). Heterogeneity was low ($I^2 = 0\%$; $P = 0.482$). Individual study results varied from a low of 2.1% transitioning (108) to a high of 3.4% (106).

Even though heterogeneity was high for the 7-day window, the small number of studies available prevented examination of the sources of this heterogeneity. However, given that the ranges of transition values were very narrow across studies and time periods (1% to 3.3%), there is confidence that repeat testing by NAAT testing within 7 days is unlikely to provide a substantial increase in the number of positive *C. difficile* results.

Level of Evidence across Questions

Table 11 depicts the LMBP strength of body of evidence for each testing practice assessed. It represents the final level of LMBP qualitative synthesis and is based on the LMBP criteria presented in Table 3 applied to the study-level information summarized in Tables 4 and 5 as well as in Fig. 4 to 6.

While some risks of bias were identified (specifically within “patient selection” and “reference standard” criteria), the assessment of the expert panel was that these did not pose a serious threat to our confidence in the findings. In sum, the decision of the expert panel was that the quality of the evidence for all questions was good. Addi-

tionally, based on evaluations of the likelihood ratio scatter matrices (Fig. 4 to 6), the effect sizes were determined to be substantial for NAAT only, substantial for the GDH/NAAT algorithm (although uncertainty remains due to the wide confidence intervals), and moderate for the GDH/toxin/NAAT algorithm (Table 11). For repeat testing using NAAT, the effect size was minimal.

ADDITIONAL CONSIDERATIONS

Applicability and Generalizability

C. difficile testing utilizing NAAT-based algorithms provides diagnostically accurate detection of the *C. difficile* organism compared to the reference standard. *C. difficile* testing by NAAT should also not be repeated to increase diagnostic sensitivity after an initial negative test, since it is already a diagnostically sensitive test.

In addition, in Table 10, it is demonstrated that there was a statistically significantly decreased sensitivity when a stool criterion (such as diarrhea or unformed stool) was included in the studies. This implies that preanalytic variables affect the diagnostic accuracy of *C. difficile* testing. This systematic review was unable to answer the question about which diagnostic test is most accurate to make the diagnosis of *C. difficile* infection due to the fact that the large majority of studies do not include clinical outcomes. Therefore, this systematic review was focused on an intermediate outcome (Fig. 1) when NAAT is part of laboratory testing.

These findings indicate the need for improvement in reporting *C. difficile* diagnostic accuracy study results rather than genuine flaws in the research. Incomplete reporting in these peer-reviewed articles can affect the usefulness of systematic review findings. While information that is relevant to a particular diagnostic accuracy evidence base can vary (e.g., preanalytic testing criteria), the Standards for Reporting of Diagnostic Accuracy Studies (STARD) provides a list of minimum essential reporting items (113). Currently, the STARD includes a requirement for the description of eligible participants but does not specifically discuss preanalytic considerations for the samples *per se*. The influence of preanalytic factors on test performance may be better established through fuller reporting of these factors in the primary evidence base and may be considered useful for augmentation of STARD reporting standards or for ensuring that the specific preanalytical considerations fall under eligible participants.

It was determined by the expert panel that while concerns remain about how well studies report patient selection criteria, this was unlikely to compromise confidence in the results. The sensitivity analysis (Table 8) demonstrated that while the reference standards of TC and CCNA are different tests, both can be used as a reference standard without concern of decreasing diagnostic accuracy. However, studies were excluded if neither of these reference standards was used or if they were used only on discordant results. Not all studies defined the reference standard as either TC or CCNA (or both) but rather combined the outcomes of these reference standards along with other diagnostic tests to create a panel of tests to serve as the reference standard. The judgment of the expert analysts was that while the use of a panel of tests as a reference method rather than TC or CCNA alone would affect diagnostic accuracy measures, these panels were unlikely to be less accurate, and so these studies could be included in the analysis and did not compromise our confidence in the findings. However, for a study to be included in this analysis, at least TC or CCNA had to be in the panel of tests and not used only for analysis of discrepant results.

Additionally, the current evidence base on the effectiveness of *C. difficile* algorithms did not permit direct assessment of health benefits, whether direct health outcomes or surrogate outcomes as specified in the analytic framework (e.g., delay to treatment or delay to isolation) and other outcomes deemed relevant (e.g., LOS or intensive care unit [ICU] stay). Therefore, the clinical utility (i.e., the degree to which the use of a test is associated with improved health outcomes) of the algorithms examined remains unclear, although treatment and clinical care options resulting from the test information are well characterized.

Comparison to Recent Clinical Guidelines

The recent IDSA/SHEA guidelines (11) outline their recommendations for practice and are based on use of Grading of Recommendations Assessment, Development, and Evaluation (GRADE) criteria. GRADE criteria differ from LMBP criteria in their assessment of an evidence base to derive practice recommendations. GRADE criteria drive a more direct accounting of the clinical factors related to *C. difficile* testing, and as was shown in the IDSA/SHEA guidelines, there were few studies in the literature that provided clinical outcome data. Therefore, the IDSA/SHEA guidelines report limited evidence to support testing practices because the data themselves are limited. However, the guidelines elegantly interpret their findings in the setting of the clinical decision-making process and guide their recommendations based on the clinical context (11).

In this ASM-led systematic review, practice recommendations relate to detection of the toxin or toxin gene of the organism, with a focus on relevant diagnostic accuracy measures rather than the diagnosis of CDI, which (as discussed above) is based on a combination of clinical presentation and laboratory testing. Therefore, in following the LMBP framework, this systematic review sought to assess diagnostic accuracy (an intermediate outcome in health care) as the review's primary outcome of interest. In other words, this systematic review did not downgrade such evidence for being indirect to patient outcomes. While this review focused on identification of the cell wall enzyme, toxin, or toxin gene of the organism as the basis of guidance on diagnostic testing strategies, it is clear that diagnostic testing directly supports decisions about whether or not to treat. Furthermore, it is recommended that the preanalytic aspects of patient presentation should be taken into account with the interpretation of the test result.

Feasibility of Implementation

Since many of the studies do not include preanalytic variables, and given the results of Table 10, testing of formed stool may lead to overdiagnosis with a sensitive test. Clinicians should utilize clinically agreed-upon symptoms, appropriate diarrheal history (at least 3 unformed bowel movements within 24 h), and antimicrobial use history as well as exclude patients on laxatives and promotility drugs (114). Prior to changing algorithms, laboratories should base their testing decision on published data, and collect and analyze the data at their institution, in order to support the new testing practices. Assays that have high sensitivity can be utilized as long as clinicians ordering the test understand the limitations for a patient who does not meet the testing criteria (11).

As emphasized in the IDSA/SHEA guidelines (11), the aspect of health care provider education on the use and interpretation of laboratory testing needs to be critically placed in the clinical context of the patient. There are initiatives that have been implemented to assist health care providers by creating criteria by which to appropriately order laboratory testing for *C. difficile*. Some health care systems have embarked on the use of a form in the electronic medical record that the provider must fill out in order for the test to be performed, based on certain preanalytic requirements, such as frequency and consistency of bowel movements. Empowerment of the health care workers who are collecting the stool from the patient to have discussions with clinicians about whether the stool is formed or diarrheal after visualization is also critical. In their study, Truong et al. restricted the use of *C. difficile* NAAT with the following requirements for orders: ≥ 3 unformed bowel movements over 24 h and no laxative intake during the previous 48 h. However, exceptions were made for patients admitted within the previous 24 h, for patients with a rectal or ostomy tube, and if the ordering provider called to override the rejection. This policy resulted in a significant reduction in test utilization as well as reduced oral vancomycin use in these patients (115). Truong et al. also looked at outcomes for individuals with cancelled *C. difficile* orders and found that they were not worse than those for individuals whose specimens were accepted and were *C. difficile* negative (115). Quan et al. utilized automated verification at the time of computerized provider order entry to enforce appropriate CDI testing. The criteria included (i) diarrhea (≥ 3 liquid/watery stools in 24 h), (ii) no reasonable alternate cause

for diarrhea, (iii) no laxative use within 24 h, (iv) no previous CDI test result within 7 days, and (v) age of >1 year. This criterion-based testing protocol reduced testing by two-thirds and decreased rates of *C. difficile* without changing the methodology of testing (114). In addition to the published data about the policies, “gray” (unpublished) data were obtained from institutions during the call for unpublished data for this systematic review (28). Some hospitals differentiated recommendations for individuals who were in the hospital for less than or more than 4 days, and if the stay was <4 days, the individual would be tested if they had unexplained loose/unformed stools. If the stay was >4 days, the algorithm included ≥ 3 liquid/watery stools in 24 h, discontinuation of laxatives, and clinical signs/symptoms of *C. difficile* infection or epidemiological risk factors for *C. difficile*. There are numerous examples in the literature which have demonstrated that improved education and adherence to preanalytic criteria prior to testing lead to appropriate clinical utility despite the *C. difficile* test that is used.

Limitations

A major limitation of this systematic review is that a main NAAT algorithm (NAAT followed by toxin testing) is not included due to the time scope of this study. This will be included in the future update of this systematic review. One limitation is the new use of the likelihood ratio within the LMBP method, and QUADAS-2. A justification for these methods has been published (36). A significant limitation of the evidence base was the failure to incorporate preanalytic parameters and clinical outcomes in the study design. This was very common in all of the literature that was evaluated and speaks to the fact that there should be increased emphasis placed on the context of laboratory testing as well as its diagnostic accuracy. This is especially clear in the IDSA/SHEA guidelines, where a small number of articles met criteria to be referenced for diagnostic testing recommendations (11). Given this limitation, there is currently little evidence base to assess the impact of overall testing practices on population health outcomes despite a high number of studies regarding testing for *C. difficile*. Reporting for health care facility onset does take into account the type of testing that is used by the facilities, so that comparisons for testing prevalence are consistent. In addition, some molecular stool testing also includes the sample being placed in liquid medium, and this will limit the ability to assess whether the stool does not conform to the shape of the container. For these types of assays, it will be difficult to determine if the preanalytical characteristics (eligibility of the sample) have been met, leading to further confusion as to what the result means for the specific patient.

FUTURE RESEARCH

After the evaluation of a large number of studies in this systematic review, the paucity of clinical data that are collected during diagnostic accuracy studies was striking. In general, the diagnostic laboratory community should consider, in addition to diagnostic comparison studies, the inclusion of preanalytic factors and postanalytic clinical outcomes in these studies. It may be also unclear as to whether the reference standard is indeed a good representation of active toxigenic infection in a patient. If an isolate of *C. difficile* recovered from stool produces toxin in an *in vitro* test, such as growing the organism in an enrichment broth, it does not tell us that toxin is being produced in the patient. The presence of the organism does not tell us about its *in vivo* activity. The dilemma that has been created by the development of NAAT as a detection method pulls us away from the *in vivo* toxin activity by just detecting the presence of the organism toxin gene in the stool. There are ultrasensitive toxin tests (116) that are being developed to be able to increase the sensitivity of toxin testing without the oversensitivity of the detection of the toxin gene in a NAAT. Due to financial pressure being placed on hospitals which have high *C. difficile* infection rates by federal pay-for-performance programs, there are clinical microbiology laboratories that are considering changing from NAAT-only or NAAT-containing algorithms to toxin tests alone (14). These tests have been shown to be less sensitive than the three testing strategies recommended here (14). The use of insensitive toxin tests may result in an

TABLE 12 Summary of ASM practice recommendations for *C. difficile* testing

Practice category	Practice recommendation
NAAT only	Use of NAAT-only testing is recommended as a best practice for the detection of the <i>C. difficile</i> toxin gene
GDH/NAAT algorithm	Use of a GDH/NAAT algorithm is recommended as a best practice for the detection of the <i>C. difficile</i> organism/toxin gene
GDH/toxin/NAAT algorithm	Use of a GDH/toxin/NAAT algorithm is recommended as a best practice for the detection of the <i>C. difficile</i> organism, toxin, or toxin gene
Repeated testing using NAAT	A recommendation for or against repeated testing for <i>C. difficile</i> using a NAAT as a best practice cannot be made due to insufficient evidence

increased number of missed diagnoses, leading to poorer clinical outcomes; however, studies regarding these outcomes need to be performed in order to determine if there are poorer clinical outcomes. Outcome studies in which different testing approaches are compared are essential to determine optimal testing strategies (see Appendix SB in the supplemental material). Additional studies are needed on the yield and duration of repeat testing using NAAT-containing testing strategies. The overlay of clinical outcome on studies of repeat testing will be important for health care facility policies for acceptance of patient samples at certain time intervals.

CONCLUSIONS

Practice Recommendations

Recommendations are categorized as “recommended,” “not recommended,” and “no recommendation for or against due to insufficient evidence.” Recommendation categorization in this review is a function of the currently available evidence base and of the CDC LMBP method, including *a priori* analysis criteria (e.g., selected effect measure rating cutoffs, the LMBP quality assessment tool, and the LMBP strength of body of evidence matrix). The approach for recommendation categorization is described in Methods above, with criteria indicated in Table 3. ASM recommendations arising from this systematic review do not serve to endorse specific NAATs; rather, they relate to the ability to choose for each individual health care system the most appropriate *C. difficile* laboratory diagnostic test algorithm that best supports the practices of the institution (11).

Practice recommendations are summarized in Table 12, with additional details provided in the remainder of this section.

ASM Recommendation for NAAT-Only Testing

Among patients suspected of having *Clostridioides (Clostridium) difficile* infection, NAAT-only testing is a recommended practice for detection of the *C. difficile* toxin gene. The overall strength of evidence for this practice is rated as high. The pooled effect rating for 46 studies meta-analyzed is substantial (+LR = 46.0 [95% CI, 35.7, 59.2]; –LR = 0.05 [95% CI, 0.04, 0.06]). Effects across studies were consistent.

ASM Recommendation for the GDH/NAAT Algorithm

Among patients suspected of having *Clostridioides (Clostridium) difficile* infection, a GDH/NAAT algorithm is a recommended practice for detection of the *C. difficile* organism/toxin gene. The overall strength of evidence for this practice is rated as high. The pooled effect rating for 11 studies meta-analyzed is substantial (+LR = 113.5 [95% CI, 49.9, 258.1]; –LR = 0.09 [95% CI, 0.06, 0.14]). Effects across studies were consistent.

ASM Recommendation for the GDH/toxin/NAAT Algorithm

Among patients suspected of having *Clostridioides (Clostridium) difficile* infection, an algorithm including NAAT is a recommended practice for detection of the *C. difficile* organism/toxin/toxin gene. The overall strength of evidence for this practice is rated as moderate (more than 3 studies with good to moderate quality-to-effect pairings achieved a “moderate” strength-of-evidence rating). The pooled effect rating for 11 studies meta-analyzed is moderate (+LR = 155.8 [95% CI, 57.7, 420.2]; –LR = 0.11 [95% CI, 0.08, 0.16]). Effects across studies were consistent.

ASM Recommendation for Repeated Testing Using NAAT

Among patients suspected of having *Clostridioides (Clostridium) difficile* infection, due to insufficient evidence, there is no recommendation for or against repeated testing by NAAT only within 7 days when the result is negative. The overall strength of evidence for this practice is rated as insufficient. The pooled effect rating for 5 studies meta-analyzed is minimal (3% conversion from negative to positive [95% CI, 0.023 to 0.038]). Effects across studies were consistent. A limited but consistent body of generally high-quality evidence indicates that repeat testing using NAAT has minimal additional benefit for detecting the presence of the *C. difficile* toxin gene.

However, while the use of the LMBP strength of body of evidence criteria sustains an “insufficient” strength-of-evidence categorization, a “minimal” effect for diagnostic yield should be interpreted to mean that repeat testing using NAAT does not appreciably contribute to patient diagnosis of CDI. Therefore, in this context, a minimal-effect finding (when combined with a good-quality evidence base) may also be interpreted as strong evidence against the use of repeat testing by NAAT. In using the LMBP method, however, a category of “recommendation against” may be achievable when outcomes for repeat testing include outcomes such as cost of testing and time to treatment, etc. In short, repeat testing by NAAT is likely a practice to be recommended against, a finding which may be more definitively sustained by future studies. See Appendix SD in the supplemental material for guidance for future studies.

APPENDIX

GLOSSARY

analytical sensitivity The ability of a method to detect a low concentration of an analyte.

clinical outcomes Measurable changes in health, function, or quality of life that result from clinical/medical care.

clinical validity The accuracy of detection of the presence or absence of a phenotype/disease.

diagnostic odds ratio The odds of a positive test result for those with disease relative to the odds of a positive test result for those without the disease, as can be mathematically represented by dividing the positive likelihood ratio by the negative likelihood ratio. It provides a global measure of a test’s diagnostic accuracy and can assist in comparison of diagnostic accuracies between two or more index tests. However, as an overall measure of diagnostic accuracy, it obscures the clinically important trade-off between rates of false-positive and false-negative results.

diagnostic sensitivity The proportion of individuals correctly classified by the index test as having a disease, target condition, or gene of the infectious agent of interest.

diagnostic specificity The proportion of individuals correctly classified by the index test as not having a disease, target condition, or gene of the infectious agent of interest.

effect size The association between two or more studies’ outcome measures for the group in which the intervention/practice was evaluated and those for its control or comparison group. Effect size ratings can be numeric or reflect the magnitude of effect in qualitative terms. Numeric representation can be converted to qualitative values through expert consensus on cutoffs representing a “substantial,” “moderate,” or “minimal” effect; however, setting cutoffs for qualitative ratings invariably has an element of subjective judgment.

evidence summary tables Used in systematic reviews to summarize the study findings, including but not limited to study design, author, statistical summary, quality of study, magnitude of benefit, absolute risk reduction, and number needed to treat. The content of the tables depends on the topic of the review.

I^2 Statistic describing the percentage of variation across studies that is due to heterogeneity rather than chance.

interclass correlation coefficient Measures a relation between two variables of different classes, in this case between different studies.

likelihood ratio scatter matrix Used in this context to plot positive and negative likelihood ratio pairings for diagnostic test accuracy studies. The pairings are then graphed on one of four quadrants derived from established thresholds for test clinical validity.

meta-analysis The process of using statistical methods to standardize and quantitatively combine the results of similar studies in an attempt to allow inferences to be made from a collection of studies. It allows for estimates of effects across studies.

preanalytical The testing phase that occurs first in the laboratory process and may include specimen handling issues that occur even prior to the time when the specimen is received in the laboratory.

quality improvement A systematic, formal approach to the analysis of practice performance and efforts to improve performance.

reference standard/method The test, combination of tests, or procedure that is considered the best available method of categorizing participants in a study of diagnostic test accuracy as having or not having a target condition.

systematic review A scientific investigation that focuses on a specific question and uses explicit, planned scientific methods to identify, select, assess, and summarize the findings of similar but separate studies. It may or may not include a quantitative synthesis of the results from separate studies (meta-analysis).

testing algorithm Diagnostic testing that uses multiple tests in a sequence of specified actions.

test clinical utility The extent to which a test is usefully informative by contributing to improved clinical management or patient-related outcomes. Demonstrations of a test's utility have been achieved through randomized controlled trials of test-and-treat (or test-and-clinical response) interventions and through decision analysis modeling in which the probabilities of relevant outcomes are estimated.

SUPPLEMENTAL MATERIAL

Supplemental material for this article may be found at <https://doi.org/10.1128/CMR.00032-18>.

SUPPLEMENTAL FILE 1, PDF file, 0.6 MB.

SUPPLEMENTAL FILE 2, PDF file, 0.1 MB.

ACKNOWLEDGMENTS

The findings and conclusions in this article are those of the authors and do not necessarily represent the official position of the Centers for Disease Control and Prevention.

We thank the LMBP Work Group, L. Clifford McDonald, Dale Gerding, and Catherine Duff as a patient advocate.

This work received funding from the American Society for Microbiology (5U47OE000055).

Vickie Baselski, Susan Benson, Cassiana E. Bittencourt, April M. Bobenchik, Nancy E. Cornish, Jennifer Dien Bard, Peter Gilligan, Jonathan C. Gullett, Colleen S. Kraft, Joseph D. Lutgring, Thomas J. Kirn, Irving Nachamkin, J. Scott Parrott, Matthew L. Rubinstein, Robert L. Sautter, and James W. Snyder have no conflicts to declare. Monika Fischer has received honoraria from Finch Therapeutics for consulting. This company is involved in development of diagnostics targeted at the human microbiome, which has relevance to *C. difficile* diagnosis and treatment. Romney M. Humphries has received speaking and consulting honoraria from Meridian, Cepheid, and Nanosphere. He also serves as CSO of a diagnostics company. Elizabeth M. Marlowe, over the course of this meta-analysis, became an employee of a manufacturer of a C.Diff NAAT assay, but this assay was not evaluated in this systematic review. Nancy S. Miller has had recent research funding and a percentage of her salary from a diagnostic company performing *C. difficile* clinical trials. Alice S. Weissfeld has performed clinical trials for *C. difficile* testing. Since 2009, her company, Microbiology Specialists Inc., has performed toxigenic *C. difficile* cultures for clinical trials run by Cepheid, BD, Nanosphere, and IMDx. These companies used the

data to gain 510 k clearances from the FDA. The company has since worked with Diasorin Inc. (formerly Focus Diagnostics), Luminex Corporation, and GenePOC Inc. Sandra S. Richter has received research funding from Roche, bioMérieux, BD Diagnostics, Hologic, Accelerate, and Diasorin.

REFERENCES

- Schroeder LF, Robilotti E, Peterson LR, Banaei N, Dowdy DW. 2014. Economic evaluation of laboratory testing strategies for hospital-associated Clostridium difficile infection. *J Clin Microbiol* 52:489–496. <https://doi.org/10.1128/JCM.02777-13>.
- Miller BA, Chen LF, Sexton DJ, Anderson DJ. 2011. Comparison of the burdens of hospital-onset, healthcare facility-associated Clostridium difficile infection and of healthcare-associated infection due to methicillin-resistant Staphylococcus aureus in community hospitals. *Infect Control Hosp Epidemiol* 32:387–390. <https://doi.org/10.1086/659156>.
- Lessa FC, Winston LG, McDonald LC, Emerging Infections Program C. difficile Surveillance Team. 2015. Burden of Clostridium difficile infection in the United States. *N Engl J Med* 372:2369–2370. <https://doi.org/10.1056/NEJMc1505190>.
- Magee G, Strauss ME, Thomas SM, Brown H, Baumer D, Broderick KC. 2015. Impact of Clostridium difficile-associated diarrhea on acute care length of stay, hospital costs, and readmission: a multicenter retrospective study of inpatients, 2009–2011. *Am J Infect Control* 43:1148–1153. <https://doi.org/10.1016/j.ajic.2015.06.004>.
- Gao T, He B, Pan Y, Deng Q, Sun H, Liu X, Chen J, Wang S, Xia Y. 2015. Association of Clostridium difficile infection in hospital mortality: a systematic review and meta-analysis. *Am J Infect Control* 43:1316–1320. <https://doi.org/10.1016/j.ajic.2015.04.209>.
- Zhang D, Prabhu VS, Marcella SW. 2018. Attributable healthcare resource utilization and costs for patients with primary and recurrent Clostridium difficile infection in the United States. *Clin Infect Dis* 66:1326–1332. <https://doi.org/10.1093/cid/cix1021>.
- Desai K, Gupta SB, Dubberke ER, Prabhu VS, Browne C, Mast TC. 2016. Epidemiological and economic burden of Clostridium difficile in the United States: estimates from a modeling approach. *BMC Infect Dis* 16:303. <https://doi.org/10.1186/s12879-016-1610-3>.
- Kyne L, Hamel MB, Polavaram R, Kelly CP. 2002. Health care costs and mortality associated with nosocomial diarrhea due to Clostridium difficile. *Clin Infect Dis* 34:346–353. <https://doi.org/10.1086/338260>.
- McGlone SM, Bailey RR, Zimmer SM, Popovich MJ, Tian Y, Uffberg P, Muder RR, Lee BY. 2012. The economic burden of Clostridium difficile. *Clin Microbiol Infect* 18:282–289. <https://doi.org/10.1111/j.1469-0691.2011.03571.x>.
- Zimlichman E, Henderson D, Tamir O, Franz C, Song P, Yamin CK, Keohane C, Denham CR, Bates DW. 2013. Health care-associated infections: a meta-analysis of costs and financial impact on the US health care system. *JAMA Intern Med* 173:2039–2046. <https://doi.org/10.1001/jamainternmed.2013.9763>.
- McDonald LC, Gerding DN, Johnson S, Bakken JS, Carroll KC, Coffin SE, Dubberke ER, Garey KW, Gould CV, Kelly C, Loo V, Shaklee Sammons J, Sandora TJ, Wilcox MH. 2018. Clinical practice guidelines for Clostridium difficile infection in adults and children: 2017 update by the Infectious Diseases Society of America (IDSA) and Society for Healthcare Epidemiology of America (SHEA). *Clin Infect Dis* 66:e1–e48. <https://doi.org/10.1093/cid/cix1085>.
- National Academies of Sciences, Engineering, and Medicine. 2015. Improving diagnosis in health care. National Academies Press, Washington, DC.
- Newman-Toker DE. 2014. A unified conceptual model for diagnostic errors: underdiagnosis, overdiagnosis, and misdiagnosis. *Diagnosis (Berl)* 1:43–48. <https://doi.org/10.1515/dx-2013-0027>.
- Fang FC, Polage CR, Wilcox MH. 2017. Point-counterpoint: what is the optimal approach for detection of Clostridium difficile infection? *J Clin Microbiol* 55:670–680. <https://doi.org/10.1128/JCM.02463-16>.
- Crobach MJT, Baktash A, Duzsenko N, Kuijper EJ. 2018. Diagnostic guidance for C. difficile infections. *Adv Exp Med Biol* 1050:27–44. https://doi.org/10.1007/978-3-319-72799-8_3.
- Peng Z, Ling L, Stratton CW, Li C, Polage CR, Wu B, Tang YW. 2018. Advances in the diagnosis and treatment of Clostridium difficile infections. *Emerg Microbes Infect* 7:15. <https://doi.org/10.1038/s41426-017-0019-4>.
- Crobach MJ, Planche T, Eckert C, Barbut F, Terveer EM, Dekkers OM, Wilcox MH, Kuijper EJ. 2016. European Society of Clinical Microbiology and Infectious Diseases: update of the diagnostic guidance document for Clostridium difficile infection. *Clin Microbiol Infect* 22(Suppl 4):S63–S81. <https://doi.org/10.1016/j.cmi.2016.03.010>.
- Planche T, Wilcox M. 2011. Reference assays for Clostridium difficile infection: one or two gold standards? *J Clin Pathol* 64:1–5. <https://doi.org/10.1136/jcp.2010.080135>.
- Boyanton BL, Jr, Sural P, Loomis CR, Pesta C, Gonzalez-Krellwitz L, Robinson-Dunn B, Riska P. 2012. Loop-mediated isothermal amplification compared to real-time PCR and enzyme immunoassay for toxigenic Clostridium difficile detection. *J Clin Microbiol* 50:640–645. <https://doi.org/10.1128/JCM.01014-11>.
- Burnham CA, Carroll KC. 2013. Diagnosis of Clostridium difficile infection: an ongoing conundrum for clinicians and for clinical laboratories. *Clin Microbiol Rev* 26:604–630. <https://doi.org/10.1128/CMR.00016-13>.
- Petersen C, Round JL. 2014. Defining dysbiosis and its influence on host immunity and disease. *Cell Microbiol* 16:1024–1033. <https://doi.org/10.1111/cmi.12308>.
- Bagdasarian N, Rao K, Malani PN. 2015. Diagnosis and treatment of Clostridium difficile in adults: a systematic review. *JAMA* 313:398–408. <https://doi.org/10.1001/jama.2014.17103>.
- Schmidt ML, Gilligan PH. 2009. Clostridium difficile testing algorithms: what is practical and feasible? *Anaerobe* 15:270–273. <https://doi.org/10.1016/j.anaerobe.2009.10.005>.
- Peterson LR, Manson RU, Paule SM, Hacek DM, Robicsek A, Thomson RB, Jr, Kaul KL. 2007. Detection of toxigenic Clostridium difficile in stool samples by real-time polymerase chain reaction for the diagnosis of C. difficile-associated diarrhea. *Clin Infect Dis* 45:1152–1160. <https://doi.org/10.1086/522185>.
- Reller ME, Lema CA, Perl TM, Cai M, Ross TL, Speck KA, Carroll KC. 2007. Yield of stool culture with isolate toxin testing versus a two-step algorithm including stool toxin testing for detection of toxigenic Clostridium difficile. *J Clin Microbiol* 45:3601–3605. <https://doi.org/10.1128/JCM.01305-07>.
- Musher DM, Manhas A, Jain P, Nuila F, Waqar A, Logan N, Marino B, Graviss EA. 2007. Detection of Clostridium difficile toxin: comparison of enzyme immunoassay results with results obtained by cytotoxicity assay. *J Clin Microbiol* 45:2737–2739. <https://doi.org/10.1128/JCM.00686-07>.
- O'Horo JC, Jones A, Sterne M, Harper C, Safdar N. 2012. Molecular techniques for diagnosis of Clostridium difficile infection: systematic review and meta-analysis. *Mayo Clin Proc* 87:643–651. <https://doi.org/10.1016/j.mayocp.2012.02.024>.
- Christenson RH, Snyder SR, Shaw CS, Derzon JH, Black RS, Mass D, Epner P, Favoretto AM, Liebow EB. 2011. Laboratory medicine best practices: systematic evidence review and evaluation methods for quality improvement. *Clin Chem* 57:816–825. <https://doi.org/10.1373/clinchem.2010.157131>.
- Snyder SR, Favoretto AM, Baetz RA, Derzon JH, Madison BM, Mass D, Shaw CS, Layfield CD, Christenson RH, Liebow EB. 2012. Effectiveness of practices to reduce blood culture contamination: a Laboratory Medicine Best Practices systematic review and meta-analysis. *Clin Biochem* 45:999–1011. <https://doi.org/10.1016/j.clinbiochem.2012.06.007>.
- Heyer NJ, Derzon JH, Wings L, Shaw C, Mass D, Snyder SR, Epner P, Nichols JH, Gayken JA, Ernst D, Liebow EB. 2012. Effectiveness of practices to reduce blood sample hemolysis in EDs: a laboratory medicine best practices systematic review and meta-analysis. *Clin Biochem* 45:1012–1032. <https://doi.org/10.1016/j.clinbiochem.2012.08.002>.
- LaRocco MT, Franek J, Leibach EK, Weissfeld AS, Kraft CS, Sautter RL, Baselski V, Rodahl D, Peterson EJ, Cornish NE. 2016. Effectiveness of preanalytic practices on contamination and diagnostic accuracy of

- urine cultures: a Laboratory Medicine Best Practices systematic review and meta-analysis. *Clin Microbiol Rev* 29:105–147. <https://doi.org/10.1128/CMR.00030-15>.
32. Buehler SS, Madison B, Snyder SR, Derzon JH, Cornish NE, Saubolle MA, Weissfeld AS, Weinstein MP, Liebow EB, Wolk DM. 2016. Effectiveness of practices to increase timeliness of providing targeted therapy for inpatients with bloodstream infections: a Laboratory Medicine Best Practices systematic review and meta-analysis. *Clin Microbiol Rev* 29: 59–103. <https://doi.org/10.1128/CMR.00053-14>.
 33. Rubinstein M, Hirsch R, Bandyopadhyay K, Madison B, Taylor T, Ranne A, Linville M, Donaldson K, Lacbawan F, Cornish N. 2018. Effectiveness of practices to support appropriate laboratory test utilization: a Laboratory Medicine Best Practices systematic review and meta-analysis. *Am J Clin Pathol* 149:197–221. <https://doi.org/10.1093/ajcp/axq147>.
 34. Whiting PF, Rutjes AW, Westwood ME, Mallett S, Deeks JJ, Reitsma JB, Leeflang MM, Sterne JA, Bossuyt PM, QUADAS-2 Group. 2011. QUADAS-2: a revised tool for the quality assessment of diagnostic accuracy studies. *Ann Intern Med* 155:529–536. <https://doi.org/10.7326/0003-4819-155-8-201110180-00009>.
 35. Deeks JJ, Altman DG. 2004. Diagnostic tests 4: likelihood ratios. *BMJ* 329:168–169. <https://doi.org/10.1136/bmj.329.7458.168>.
 36. Rubinstein ML, Kraft CS, Parrott JS. 2018. Determining qualitative effect size ratings using a likelihood ratio scatter matrix in diagnostic test accuracy systematic reviews. *Diagnosis (Berl)* 5:205–214. <https://doi.org/10.1515/dx-2018-0061>.
 37. Rutter CM, Gatsonis CA. 2001. A hierarchical regression approach to meta-analysis of diagnostic test accuracy evaluations. *Stat Med* 20: 2865–2884. <https://doi.org/10.1002/sim.942>.
 38. Harbord RM, Deeks JJ, Egger M, Whiting P, Sterne JA. 2007. A unification of models for meta-analysis of diagnostic accuracy studies. *Biostatistics* 8:239–251. <https://doi.org/10.1093/biostatistics/kxl004>.
 39. Alcalá L, Reigadas E, Marin M, Fernandez-Chico A, Catalan P, Bouza E. 2015. Comparison of GenomEra C. difficile and Xpert C. difficile as confirmatory tests in a multistep algorithm for diagnosis of *Clostridium difficile* infection. *J Clin Microbiol* 53:332–335. <https://doi.org/10.1128/JCM.03093-14>.
 40. Babady NE, Stiles J, Ruggiero P, Khosa P, Huang D, Shuptrav S, Kamboj M, Kiehn TE. 2010. Evaluation of the Cepheid Xpert *Clostridium difficile* Epi assay for diagnosis of *Clostridium difficile* infection and typing of the NAP1 strain at a cancer hospital. *J Clin Microbiol* 48:4519–4524. <https://doi.org/10.1128/JCM.01648-10>.
 41. Barbut F, Braun M, Burghoffer B, Lalande V, Eckert C. 2009. Rapid detection of toxigenic strains of *Clostridium difficile* in diarrheal stools by real-time PCR. *J Clin Microbiol* 47:1276–1277. <https://doi.org/10.1128/JCM.00309-09>.
 42. Barbut F, Monot M, Rousseau A, Cavelot S, Simon T, Burghoffer B, Lalande V, Tankovic J, Petit JC, Dupuy B, Eckert C. 2011. Rapid diagnosis of *Clostridium difficile* infection by multiplex real-time PCR. *Eur J Clin Microbiol Infect Dis* 30:1279–1285. <https://doi.org/10.1007/s10096-011-1224-z>.
 43. Barkin JA, Nandi N, Miller N, Grace A, Barkin JS, Sussman DA. 2012. Superiority of the DNA amplification assay for the diagnosis of *C. difficile* infection: a clinical comparison of fecal tests. *Dig Dis Sci* 57:2592–2599. <https://doi.org/10.1007/s10620-012-2200-x>.
 44. Beck ET, Buchan BW, Riebe KM, Alkins BR, Pancholi P, Granato PA, Ledebner NA. 2014. Multicenter evaluation of the Quidel Lyra Direct C. difficile nucleic acid amplification assay. *J Clin Microbiol* 52:1998–2002. <https://doi.org/10.1128/JCM.03089-13>.
 45. Brown NA, Lebar WD, Young CL, Hankerd RE, Newton DW. 2011. Diagnosis of *Clostridium difficile* infection: comparison of four methods on specimens collected in Cary-Blair transport medium and tcdB PCR on fresh versus frozen samples. *Infect Dis Rep* 3:e5. <https://doi.org/10.4081/idr.2011.e5>.
 46. Bruins MJ, Verbeek E, Wallinga JA, Bruijnesteijn van Coppenraet LE, Kuijper EJ, Bloembergen P. 2012. Evaluation of three enzyme immunoassays and a loop-mediated isothermal amplification test for the laboratory diagnosis of *Clostridium difficile* infection. *Eur J Clin Microbiol Infect Dis* 31:3035–3039. <https://doi.org/10.1007/s10096-012-1658-y>.
 47. Buchan BW, Mackey TL, Daly JA, Alger G, Denys GA, Peterson LR, Kehl SC, Ledebner NA. 2012. Multicenter clinical evaluation of the Portrait toxigenic C. difficile assay for detection of toxigenic *Clostridium difficile* strains in clinical stool specimens. *J Clin Microbiol* 50:3932–3936. <https://doi.org/10.1128/JCM.02083-12>.
 48. Calderaro A, Buttrini M, Martinelli M, Gorrini C, Montecchini S, Medici MC, Arcangeletti MC, De Conto F, Covan S, Chezzi C. 2013. Comparative analysis of different methods to detect *Clostridium difficile* infection. *New Microbiol* 36:57–63.
 49. Carroll KC, Buchan BW, Tan S, Stamper PD, Riebe KM, Pancholi P, Kelly C, Rao A, Fader R, Cavagnolo R, Watson W, Goering RV, Trevino EA, Weissfeld AS, Ledebner NA. 2013. Multicenter evaluation of the Verigene *Clostridium difficile* nucleic acid assay. *J Clin Microbiol* 51: 4120–4125. <https://doi.org/10.1128/JCM.01690-13>.
 50. de Boer RF, Wijma JJ, Schuurman T, Moedt J, Dijk-Alberts BG, Ott A, Kooistra-Smid AM, van Duynhoven YT. 2010. Evaluation of a rapid molecular screening approach for the detection of toxigenic *Clostridium difficile* in general and subsequent identification of the tcdC delta117 mutation in human stools. *J Microbiol Methods* 83:59–65. <https://doi.org/10.1016/j.mimet.2010.07.017>.
 51. de Jong E, de Jong AS, Bartels CJ, van der Rijt-van den Biggelaar C, Melchers WJ, Sturm PD. 2012. Clinical and laboratory evaluation of a real-time PCR for *Clostridium difficile* toxin A and B genes. *Eur J Clin Microbiol Infect Dis* 31:2219–2225. <https://doi.org/10.1007/s10096-012-1558-1>.
 52. Devlin HR, Jackson C, Pangan O, Broukhanski G, Pillai DR. 2011. Rapid and accurate diagnosis of *Clostridium difficile* infection. AMMI Canada–CACMID Annual Conference, poster P14. *Can J Infect Dis Med Microbiol* 22(Suppl A):24A. <https://doi.org/10.1155/2011/760460>.
 53. Doing KM, Hintz MS. 2012. Prospective evaluation of the Meridian Illumigene loop-mediated amplification assay and the Gen Probe Pro-Gastro Cd polymerase chain reaction assay for the direct detection of toxigenic *Clostridium difficile* from fecal samples. *Diagn Microbiol Infect Dis* 72:8–13. <https://doi.org/10.1016/j.diagmicrobio.2011.09.008>.
 54. Eastwood K, Else P, Charlett A, Wilcox M. 2009. Comparison of nine commercially available *Clostridium difficile* toxin detection assays, a real-time PCR assay for *C. difficile* tcdB, and a glutamate dehydrogenase detection assay to cytotoxin testing and cytotoxigenic culture methods. *J Clin Microbiol* 47:3211–3217. <https://doi.org/10.1128/JCM.01082-09>.
 55. Eckert C, Holscher E, Petit A, Lalande V, Barbut F. 2014. Molecular test based on isothermal helicase-dependent amplification for detection of the *Clostridium difficile* toxin A gene. *J Clin Microbiol* 52:2386–2389. <https://doi.org/10.1128/JCM.00353-14>.
 56. Floré K. 2012. A new player on the market for detection of *Clostridium difficile*? 21st ECCMID/27th ICC, Milan Italy. *Clin Microb Infect* 17(Suppl 4):S583. <https://doi.org/10.1111/j.1469-0691.2011.03558.x>.
 57. Goldenberg SD, Cliff PR, French GL. 2010. Laboratory diagnosis of *Clostridium difficile* infection. *J Clin Microbiol* 48:3048–3049. <https://doi.org/10.1128/JCM.00223-10>.
 58. Gyorke CE, Wang S, Leslie JL, Cohen SH, Solnick JV, Polage CR. 2013. Evaluation of *Clostridium difficile* fecal load and limit of detection during a prospective comparison of two molecular tests, the illumigene C. difficile and Xpert C. difficile/Epi tests. *J Clin Microbiol* 51:278–280. <https://doi.org/10.1128/JCM.02120-12>.
 59. Kim H, Jeong SH, Kim M, Lee Y, Lee K. 2012. Detection of *Clostridium difficile* toxin A/B genes by multiplex real-time PCR for the diagnosis of *C. difficile* infection. *J Med Microbiol* 61:274–277. <https://doi.org/10.1099/jmm.0.035618-0>.
 60. Hirvonen JJ, Mentula S, Kaukoranta SS. 2013. Evaluation of a new automated homogeneous PCR assay, GenomEra C. difficile, for rapid detection of toxigenic *Clostridium difficile* in fecal specimens. *J Clin Microbiol* 51:2908–2912. <https://doi.org/10.1128/JCM.01083-13>.
 61. Hoegh AM, Nielsen JB, Lester A, Friis-Møller A, Schønning K. 2012. A multiplex, internally controlled real-time PCR assay for detection of toxigenic *Clostridium difficile* and identification of hypervirulent strain 027/ST-1. *Eur J Clin Microbiol Infect Dis* 31:1073–1079. <https://doi.org/10.1007/s10096-011-1409-5>.
 62. Hong G, Park KS, Ki CS, Lee NY. 2014. Evaluation of the illumigene C. difficile assay for toxigenic *Clostridium difficile* detection: a prospective study of 302 consecutive clinical fecal samples. *Diagn Microbiol Infect Dis* 80:177–180. <https://doi.org/10.1016/j.diagmicrobio.2014.08.014>.
 63. Huang H, Weintraub A, Fang H, Nord CE. 2009. Comparison of a commercial multiplex real-time PCR to the cell cytotoxicity neutralization assay for diagnosis of *Clostridium difficile* infections. *J Clin Microbiol* 47:3729–3731. <https://doi.org/10.1128/JCM.01280-09>.
 64. Humphries RM, Uslan DZ, Rubin Z. 2013. Performance of *Clostridium difficile* toxin enzyme immunoassay and nucleic acid amplification tests stratified by patient disease severity. *J Clin Microbiol* 51:869–873. <https://doi.org/10.1128/JCM.02970-12>.

65. Jazmati N, Wiegel P, Licanin B, Plum G. 2015. Evaluation of the Qiagen artus C. difficile QS-RGQ kit for detection of *Clostridium difficile* toxins A and B in clinical stool specimens. *J Clin Microbiol* 53:1942–1944. <https://doi.org/10.1128/JCM.00613-15>.
66. Jensen MB, Olsen KE, Nielsen XC, Hoegh AM, Dessau RB, Atlung T, Engberg J. 2015. Diagnosis of *Clostridium difficile*: real-time PCR detection of toxin genes in faecal samples is more sensitive compared to toxigenic culture. *Eur J Clin Microbiol Infect Dis* 34:727–736. <https://doi.org/10.1007/s10096-014-2284-7>.
67. Johnstone D, Toye B, Desjardins M. 2010. Laboratory diagnosis of C. difficile (CD) infection (CDI): Techlab C difficile Tox A/B II EIA (TOX AB), Techlab C. difficile CHEK-60 EIA (GDH), Becton Dickinson GeneOhm Cdiff PCR (BD-PCR), and the Prodesse ProGastro CD PCRs (PG-PCR) compared to toxigenic culture. AMMI Canada-CACMID 2010 Annual Conference, Edmonton, AB, Canada, poster C2. *Can J Infect Dis Med Microbiol* 21(Suppl A):5A. <https://doi.org/10.1155/2010/723876>.
68. Karre T, Sloan L, Patel R, Mandrekar J, Rosenblatt J. 2011. Comparison of two commercial molecular assays to a laboratory-developed molecular assay for diagnosis of *Clostridium difficile* infection. *J Clin Microbiol* 49:725–727. <https://doi.org/10.1128/JCM.01028-10>.
69. Kato H, Yokoyama T, Kato H, Arakawa Y. 2005. Rapid and simple method for detecting the toxin B gene of *Clostridium difficile* in stool specimens by loop-mediated isothermal amplification. *J Clin Microbiol* 43:6108–6112. <https://doi.org/10.1128/JCM.43.12.6108-6112.2005>.
70. Kato N, Ou CY, Kato H, Bartley SL, Luo CC, Killgore GE, Ueno K. 1993. Detection of toxigenic *Clostridium difficile* in stool specimens by the polymerase chain reaction. *J Infect Dis* 167:455–458. <https://doi.org/10.1093/infdis/167.2.455>.
71. Knetsch CW, Bakker D, de Boer RF, Sanders I, Hofs S, Kooistra-Smid AM, Corver J, Eastwood K, Wilcox MH, Kuijper EJ. 2011. Comparison of real-time PCR techniques to cytotoxigenic culture methods for diagnosing *Clostridium difficile* infection. *J Clin Microbiol* 49:227–231. <https://doi.org/10.1128/JCM.01743-10>.
72. Kvach EJ, Ferguson D, Riska PF, Landry ML. 2010. Comparison of BD GeneOhm Cdiff real-time PCR assay with a two-step algorithm and a toxin A/B enzyme-linked immunosorbent assay for diagnosis of toxigenic *Clostridium difficile* infection. *J Clin Microbiol* 48:109–114. <https://doi.org/10.1128/JCM.01630-09>.
73. Kyryvenko ON, Samuel LP, Tibbetts RJ. 2010. Comparison of detection methods for the rapid identification of toxigenic *Clostridium difficile*. 99th Annual USCAP Meeting, Washington, DC, 20 to 26 March 2010, poster 1490. *Lab Invest* 90:334A.
74. Lalande V, Barrault L, Wadel S, Eckert C, Petit JC, Barbut F. 2011. Evaluation of a loop-mediated isothermal amplification assay for diagnosis of *Clostridium difficile* infections. *J Clin Microbiol* 49:2714–2716. <https://doi.org/10.1128/JCM.01835-10>.
75. Landry ML, Ferguson D, Topal J. 2014. Comparison of Simplexa universal direct PCR with cytotoxicity assay for diagnosis of *Clostridium difficile* infection: performance, cost, and correlation with disease. *J Clin Microbiol* 52:275–280. <https://doi.org/10.1128/JCM.02545-13>.
76. Langley AJ, Prime K, Burnie JP. 1995. Comparison of culture, cytotoxin assay, two enzyme-linked immunosorbent assays and the polymerase chain reaction in the laboratory diagnosis of *Clostridium difficile*-associated disease. *Serodiagn Immunother Infect* 7:135–140. [https://doi.org/10.1016/0888-0786\(95\)97898-F](https://doi.org/10.1016/0888-0786(95)97898-F).
77. Larson AM, Fung AM, Fang FC. 2010. Evaluation of tcdB real-time PCR in a three-step diagnostic algorithm for detection of toxigenic *Clostridium difficile*. *J Clin Microbiol* 48:124–130. <https://doi.org/10.1128/JCM.00734-09>.
78. Le Guern R, Herwegh S, Grandbastien B, Courcol R, Wallet F. 2012. Evaluation of a new molecular test, the BD Max Cdiff, for detection of toxigenic *Clostridium difficile* in fecal samples. *J Clin Microbiol* 50:3089–3090. <https://doi.org/10.1128/JCM.01250-12>.
79. Leitner E, Einetter M, Grisold AJ, Marth E, Feierl G. 2013. Evaluation of the BD MAX Cdiff assay for the detection of the toxin B gene of *Clostridium difficile* out of faecal specimens. *Diagn Microbiol Infect Dis* 76:390–391. <https://doi.org/10.1016/j.diagmicrobio.2013.03.007>.
80. McElgunn CJ, Pereira CR, Parham NJ, Smythe JE, Wigglesworth MJ, Smielewska A, Parmar SA, Gandelman OA, Brown NM, Tisi LC, Curran MD. 2014. A low complexity rapid molecular method for detection of *Clostridium difficile* in stool. *PLoS One* 9:e83808. <https://doi.org/10.1371/journal.pone.0083808>.
81. Miller S, Wiita A, Wright C, Reyes H, Liu C. 2013. Evaluation of glutamate dehydrogenase immunoassay screening with toxin confirmation for the diagnosis of *Clostridium difficile* infection. *Lab Med* 44:e65–e71. <https://doi.org/10.1309/LM31ZX1PFRZTGUJ>.
82. Noren T, Alriksson I, Andersson J, Akerlund T, Unemo M. 2011. Rapid and sensitive loop-mediated isothermal amplification test for *Clostridium difficile* detection challenges cytotoxin B cell test and culture as gold standard. *J Clin Microbiol* 49:710–711. <https://doi.org/10.1128/JCM.01824-10>.
83. Noren T, Unemo M, Magnusson C, Eiserman M, Matussek A. 2014. Evaluation of the rapid loop-mediated isothermal amplification assay Illumigene for diagnosis of *Clostridium difficile* in an outbreak situation. *APMIS* 122:155–160. <https://doi.org/10.1111/apm.12121>.
84. Novak-Weekley SM, Marlowe EM, Miller JM, Cumpio J, Nomura JH, Vance PH, Weissfeld A. 2010. *Clostridium difficile* testing in the clinical laboratory by use of multiple testing algorithms. *J Clin Microbiol* 48:889–893. <https://doi.org/10.1128/JCM.01801-09>.
85. Pallis A, Jazayeri J, Ward P, Dimovski K, Svobodova S. 2013. Rapid detection of *Clostridium difficile* toxins from stool samples using real-time multiplex PCR. *J Med Microbiol* 62:1350–1356. <https://doi.org/10.1099/jmm.0.058339-0>.
86. Peterson LR, Mehta MS, Patel PA, Hacek DM, Harazin M, Nagwekar PP, Thomson RB, Jr, Robicsek A. 2011. Laboratory testing for *Clostridium difficile* infection: light at the end of the tunnel. *Am J Clin Pathol* 136:372–380. <https://doi.org/10.1309/AJCPPT5XKRNSXVIL>.
87. Novatschit P, Morgan J, Bradford D, Engelhardt N, Riley TV. 2015. Evaluation of the BD Max Cdiff assay for the detection of toxigenic *Clostridium difficile* in human stool specimens. *Pathology* 47:165–168. <https://doi.org/10.1097/PAT.0000000000000214>.
88. Shin S, Kim M, Kim M, Lim H, Kim H, Lee K, Chong Y. 2012. Evaluation of the Xpert *Clostridium difficile* assay for the diagnosis of *Clostridium difficile* infection. *Ann Lab Med* 32:355–358. <https://doi.org/10.3343/alm.2012.32.5.355>.
89. Silva RO, Vilela EG, Neves MS, Lobato FC. 2014. Evaluation of three enzyme immunoassays and a nucleic acid amplification test for the diagnosis of *Clostridium difficile*-associated diarrhea at a university hospital in Brazil. *Rev Soc Bras Med Trop* 47:447–450. <https://doi.org/10.1590/0037-8682-0100-2014>.
90. Soh YS, Yang JJ, You E, La Jeon Y, Kim MJ, Nam YS, Cho SY, Park TS, Lee HJ. 2014. Comparison of two molecular methods for detecting toxigenic *Clostridium difficile*. *Ann Clin Lab Sci* 44:27–31.
91. Stamper PD, Alcabasa R, Aird D, Babier W, Wehrin J, Ikpeama I, Carroll KC. 2009. Comparison of a commercial real-time PCR assay for tcdB detection to a cell culture cytotoxicity assay and toxigenic culture for direct detection of toxin-producing *Clostridium difficile* in clinical samples. *J Clin Microbiol* 47:373–378. <https://doi.org/10.1128/JCM.01613-08>.
92. Swindells J, Brenwald N, Reading N, Oppenheim B. 2010. Evaluation of diagnostic tests for *Clostridium difficile* infection. *J Clin Microbiol* 48:606–608. <https://doi.org/10.1128/JCM.01579-09>.
93. Terhes G, Urban E, Soki J, Nacsá E, Nagy E. 2009. Comparison of a rapid molecular method, the BD GeneOhm Cdiff assay, to the most frequently used laboratory tests for detection of toxin-producing *Clostridium difficile* in diarrheal feces. *J Clin Microbiol* 47:3478–3481. <https://doi.org/10.1128/JCM.01133-09>.
94. Tojo M, Nagamatsu M, Hayakawa K, Mezaki K, Kirikae T, Ohmagari N. 2014. Evaluation of an automated rapid diagnostic test for detection of *Clostridium difficile*. *PLoS One* 9:e106102. <https://doi.org/10.1371/journal.pone.0106102>.
95. Van Broeck JD, Hubert C, Vast M, Delmee M. 2010. A two-step algorithm for the diagnosis of C. difficile infection: screening with a rapid immunoassay for the detection of glutamate dehydrogenase and toxins A and B followed by real-time PCR for C. difficile, abstr P680. *Abstr 20th ECCMID*.
96. Van Broeck JD, Delmee M. 2012. Evaluation of a loop-mediated isothermal amplification technique for detection, abstr P2259. *Abstr 22nd ECCMID*.
97. Vasoo S, Stevens J, Portillo L, Barza R, Schejbal D, Wu MM, Chancey C, Singh K. 2014. Cost-effectiveness of a modified two-step algorithm using a combined glutamate dehydrogenase/toxin enzyme immunoassay and real-time PCR for the diagnosis of *Clostridium difficile* infection. *J Microbiol Immunol Infect* 47:75–78. <https://doi.org/10.1016/j.jmii.2012.07.008>.
98. van den Berg RJ, Bruijnesteijn van Coppenraet LS, Gerritsen HJ, Endtz HP, van der Vorm ER, Kuijper EJ. 2005. Prospective multicenter evaluation of a new immunoassay and real-time PCR for rapid diagnosis of

- Clostridium difficile-associated diarrhea in hospitalized patients. *J Clin Microbiol* 43:5338–5340. <https://doi.org/10.1128/JCM.43.10.5338-5340.2005>.
99. van den Berg RJ, Kuijper EJ, van Coppenraet LE, Claas EC. 2006. Rapid diagnosis of toxinogenic *Clostridium difficile* in faecal samples with internally controlled real-time PCR. *Clin Microbiol Infect* 12:184–186. <https://doi.org/10.1111/j.1469-0691.2005.01301.x>.
 100. van den Berg RJ, Vaessen N, Endtz HP, Schulin T, van der Vorm ER, Kuijper EJ. 2007. Evaluation of real-time PCR and conventional diagnostic methods for the detection of *Clostridium difficile*-associated diarrhoea in a prospective multicentre study. *J Med Microbiol* 56: 36–42. <https://doi.org/10.1099/jmm.0.46680-0>.
 101. Viala C, Le Monnier A, Maataoui N, Rousseau C, Collignon A, Poilane I. 2012. Comparison of commercial molecular assays for toxigenic *Clostridium difficile* detection in stools: BD GeneOhm Cdiff, XPert C. difficile and illumigene C. difficile. *J Microbiol Methods* 90:83–85. <https://doi.org/10.1016/j.mimet.2012.04.017>.
 102. Walkty A, Lagace-Wiens PR, Manickam K, Adam H, Pieroni P, Hoban D, Karlowsky JA, Alfa M. 2013. Evaluation of an algorithmic approach in comparison with the Illumigene assay for laboratory diagnosis of *Clostridium difficile* infection. *J Clin Microbiol* 51:1152–1157. <https://doi.org/10.1128/JCM.03203-12>.
 103. Ylisiurua P, Koskela M, Vainio O, Tuokko H. 2013. Comparison of antigen and two molecular methods for the detection of *Clostridium difficile* toxins. *Scand J Infect Dis* 45:19–25. <https://doi.org/10.3109/00365548.2012.708780>.
 104. Zidaric V, Kevorkijan BK, Oresic N, Janezic S, Rupnik M. 2011. Comparison of two commercial molecular tests for the detection of *Clostridium difficile* in the routine diagnostic laboratory. *J Med Microbiol* 60: 1131–1136. <https://doi.org/10.1099/jmm.0.030163-0>.
 105. Green DA, Stotler B, Jackman D, Whittier S, Della-Latta P. 2014. Clinical characteristics of patients who test positive for *Clostridium difficile* by repeat PCR. *J Clin Microbiol* 52:3853–3855. <https://doi.org/10.1128/JCM.01659-14>.
 106. Khanna S, Pardi DS, Rosenblatt JE, Patel R, Kammer PP, Baddour LM. 2012. An evaluation of repeat stool testing for *Clostridium difficile* infection by polymerase chain reaction. *J Clin Gastroenterol* 46: 846–849. <https://doi.org/10.1097/MCG.0b013e3182432273>.
 107. Luo RF, Banaei N. 2010. Is repeat PCR needed for diagnosis of *Clostridium difficile* infection? *J Clin Microbiol* 48:3738–3741. <https://doi.org/10.1128/JCM.00722-10>.
 108. Nistico JA, Hage JE, Schoch PE, Cunha BA. 2013. Unnecessary repeat *Clostridium difficile* PCR testing in hospitalized adults with *C. difficile*-negative diarrhea. *Eur J Clin Microbiol Infect Dis* 32:97–99. <https://doi.org/10.1007/s10096-012-1719-2>.
 109. Deshpande A, Pasupuleti V, Patel P, Pant C, Pagadala M, Hall G, Hu B, Jain A, Rolston DD, Sferra TJ, Atreja A. 2012. Repeat stool testing for *Clostridium difficile* using enzyme immunoassay in patients with inflammatory bowel disease increases diagnostic yield. *Curr Med Res Opin* 28:1553–1560. <https://doi.org/10.1185/03007995.2012.717529>.
 110. Macaskill P, Walter SD, Irwig L, Franco EL. 2002. Assessing the gain in diagnostic performance when combining two diagnostic tests. *Stat Med* 21:2527–2546. <https://doi.org/10.1002/sim.1227>.
 111. Sutton AJ, Higgins JP. 2008. Recent developments in meta-analysis. *Stat Med* 27:625–650. <https://doi.org/10.1002/sim.2934>.
 112. Macaskill P, Gatsonis C, Deeks JJ, Harbord RM, Takwoingi Y. 2010. Chapter 10: Analysing and presenting results, p 20. *In* Deeks JJ, Bossuyt PM, Gatsonis C (ed), *Cochrane handbook for systematic reviews of diagnostic test accuracy version 10*. The Cochrane Collaboration, London, United Kingdom. <https://methods.cochrane.org/sites/methods.cochrane.org.sdt/files/public/uploads/Chapter%2010%20-%20Version%201.0.pdf>.
 113. Bossuyt PM, Reitsma JB, Bruns DE, Gatsonis CA, Glasziou PP, Irwig L, Lijmer JG, Moher D, Rennie D, de Vet HC, Kressel HY, Rifai N, Golub RM, Altman DG, Hoof L, Korevaar DA, Cohen JF, STARD Group. 2015. STARD 2015: an updated list of essential items for reporting diagnostic accuracy studies. *Clin Chem* 61:1446–1452. <https://doi.org/10.1373/clinchem.2015.246280>.
 114. Quan KA, Yim J, Merrill D, Khusbu U, Madey K, Dickey L, Dangodara AA, Rudkin SE, O'Brien M, Thompson D, Parekh N, Albers CG, Wilson WC, Thrupp L, Bittencourt CE, Huang SS, Gohil SK. 2018. Reductions in *Clostridium difficile* infection (CDI) rates using real-time automated clinical criteria verification to enforce appropriate testing. *Infect Control Hosp Epidemiol* 39:625–627. <https://doi.org/10.1017/ice.2018.32>.
 115. Truong CY, Gombar S, Wilson R, Sundararajan G, Tekic N, Holubar M, Shepard J, Madison A, Tompkins L, Shah N, Deresinski S, Schroeder LF, Banaei N. 2017. Real-time electronic tracking of diarrheal episodes and laxative therapy enables verification of *Clostridium difficile* clinical testing criteria and reduction of *Clostridium difficile* infection rates. *J Clin Microbiol* 55:1276–1284. <https://doi.org/10.1128/JCM.02319-16>.
 116. Pollock NR, Banz A, Chen X, Williams D, Xu H, Cuddemi CA, Cui AX, Perrotta M, Alhassan E, Riou B, Lantz A, Miller MA, Kelly CP. 2018. Comparison of *C. difficile* stool toxin concentrations in adults with symptomatic infection and asymptomatic carriage using an ultrasensitive quantitative immunoassay. *Clin Infect Dis* 68:78–86. <https://doi.org/10.1093/cid/ciy415>.

Colleen S. Kraft earned her M.D. from Indiana University in 2002 and her M.Sc. from Emory University in 2013. She completed a residency in internal medicine and a fellowship in infectious diseases in the Department of Medicine, Emory University. She concluded her postgraduate clinical training in an Accreditation Council for Graduate Medical Education (ACGME) fellowship in medical microbiology in the Department of Pathology and Laboratory Medicine, Emory University. She has a joint faculty appointment at Emory University in the Department of Pathology and Laboratory Medicine and the Division of Infectious Diseases. She is the medical director of the microbiology section at Emory Healthcare and program director of the medical microbiology fellowship and performs rounds for the inpatient infectious diseases service. Her interests include microbiology questions that have direct clinical implication and the support of these practices with evidence-based methods. She has been on the faculty at Emory since 2010.



J. Scott Parrott received his Ph.D. from the University of Chicago. He has over 25 years of experience working as a methodological and statistical consultant both within and outside the university setting in such areas as medical and health care research, international law, public health, evidence-based medicine, childhood obesity, urban poverty, education, and international aid. Dr. Parrott's recent work has focused on evidence analysis methods and statistical modeling of intervention and outcome and evaluation studies. For the past 8 years, he has worked as a professor at Rutgers University in New Jersey with appointments in the Departments of Nutritional Sciences and Interdisciplinary Studies and the Department of Epidemiology in the School of Public Health. He currently teaches intro and advanced statistical methods and evidence analysis to master's and doctoral students.



Nancy E. Cornish received her B.A. in philosophy and M.D. from the University of Vermont, Burlington; she then completed her pathology residency at the Vermont Medical Center Hospital. She completed a microbiology fellowship at the Cleveland Clinic Foundation before serving 13 years as a pathologist and director of microbiology at the Methodist Health Care Systems and Children's Hospital in Omaha, NE. Dr. Cornish has been active in laboratory professional organizations as a faculty member, speaker, and presenter. She has spent over 20 years with the College of American Pathologists and the American Society of Clinical Pathologists and nearly 20 years with the American Society for Microbiology and the Infectious Diseases Society of America. She holds certification with the American Board of Pathologists in anatomic and clinical pathology, with a special qualification in medical microbiology. Dr. Cornish is currently medical officer for the Division of Laboratory Programs, Standards, and Services at the Centers for Disease Control and Prevention.



Matthew L. Rubinstein is a health scientist at the Centers for Disease Control and Prevention (CDC) Division of Laboratory Systems, focused on the CDC Laboratory Medicine Best Practices (LMBP) initiative. He also is a part-time adjunct instructor at Rutgers University School of Health Professions teaching graduate students the skills of evidence analysis. He has been an ASCP-certified medical technologist since 2002 and earned B.S. degrees in Biology and Clinical Laboratory Science from the University of Texas at Austin and the University of Texas Health Science Center at San Antonio, respectively. He earned a master's of science in Health Science (clinical laboratory science track) from Rutgers University School of Health Related Professions. Prior to laboratory medicine quality improvement work at the CDC, he was technical supervisor at the Armed Services Blood Bank Center—Pacific Northwest and also served in the U.S. Army as a clinical laboratory officer.



Alice S. Weissfeld is the president, chief executive officer, and laboratory director of Microbiology Specialists Incorporated, a reference laboratory that she cofounded in 1984. She earned her doctoral degree in microbiology at Rutgers University, New Brunswick, NJ, and completed a postdoctoral fellowship in public health and medical laboratory microbiology at Baylor College of Medicine, Houston, TX. Dr. Weissfeld is a diplomate of the American Board of Medical Microbiology of the American Academy of Microbiology and was elected to a fellowship in the American Academy of Microbiology in 1986. She is currently a member of ASM's new Professional Practice Committee and chair of its Subcommittee on Evidence-Based Medicine. The Subcommittee oversees the development of laboratory medicine best practice guidelines using a robust, methodical, and transparent method developed by the CDC and is participating in a cooperative agreement with the CDC to evaluate metrics to encourage the dissemination of ASM's guidelines.



Peggy McNult earned her bachelor's degree in Political Science from the University of Maryland. She is the Director of Clinical and Public Health Microbiology at the American Society for Microbiology (ASM). She works to advance the practice of clinical and public health microbiology by working with volunteers to develop and deliver professional development content for members to increase their knowledge, enhance their skills, and contribute to the profession. As a newer governance unit, she has created several new programs, including evidence-based guidelines, while overseeing ASM's evergreen certification and accreditation programs. She has worked at ASM for more than 30 years in several different departments, allowing her the luxury to develop relationships with many members who work in multiple disciplines of the microbial sciences.



Irving Nachamkin is Professor of Pathology and Laboratory Medicine at the Perelman School of Medicine at the University of Pennsylvania and Director of Laboratory Medicine at the Hospital of the University of Pennsylvania in Philadelphia, PA. Dr. Nachamkin received his doctor and master's of Public Health degrees from the University of North Carolina at Chapel Hill. Following his doctoral education, he completed an ASM CPEP Fellowship in Public Health and Medical Microbiology at the Medical College of Virginia, Virginia Commonwealth University, Richmond, VA, prior to beginning his career at Penn in 1982. He has longstanding clinical and research interests in gastrointestinal infections, including infections with *Clostridium difficile*.



Romney M. Humphries is the Chief Scientific Officer for Accelerate Diagnostics. Prior to this, she served as Section Chief of Clinical Microbiology at UCLA. She received her Ph.D. in bacterial pathogenesis from University of Calgary, Alberta, Canada, and completed a CPEP fellowship in Clinical and Public Health Microbiology at UCLA. She is a member of the CLSI AST Subcommittee, the CAP Microbiology Resource Committee, and the American Society for Microbiology Clinical Laboratory Practices Committee. Her research interests are antimicrobial resistance testing.



Continued next page

Thomas J. Kirn, M.D., Ph.D., F.A.C.P., D.(A.B.M.M.), received his M.D. and Ph.D. in Medicine and in Microbiology and Immunology from the Geisel School of Medicine at Dartmouth and then trained in Pathology and Laboratory Medicine at the Hospital of the University of Pennsylvania. He holds a faculty position at Rutgers Robert Wood Johnson Medical School as Associate Professor of Pathology and Laboratory Medicine and Medicine (Infectious Diseases). He is a board-certified Clinical Pathologist and Microbiologist who currently serves as the Director of Infectious Disease Diagnostics and Molecular Pathology at Robert Wood Johnson University Hospital (RWJUH). He is also the Medical Director of the New Jersey Public Health Laboratory. His present research interests include evaluation of the impact of infectious disease diagnostics on clinical outcomes, laboratory stewardship, and point-of-care infectious disease diagnostics. He has been working in the Infectious Disease Diagnostics field for 10 years.



Jennifer Dien Bard earned her Ph.D. in Medical Sciences at the University of Alberta, Canada, followed by completion of a 2-year postdoctoral fellowship in the Medical and Public Health Microbiology at UCLA. She started her career in 2010 as a Clinical Microbiologist and Assistant Professor at Queen's University in Kingston, Ontario. In 2012, Dr. Dien Bard took on the role of Director of Clinical Microbiology and Virology at Children's Hospital Los Angeles. She is also an Associate Professor of Clinical Pathology in the Department of Pathology, Keck School of Medicine, University of Southern California. She is a Diplomate of the American Board of Medical Microbiology and a Fellow of the Canadian College of Microbiologists. Dr. Dien Bard has 13 years of Clinical Microbiology experience and has witness firsthand the importance of developing best practices, particularly in the era of novel and innovative technologies.



Joseph D. Lutgring completed his B.S. in Biology from Indiana University, Bloomington, prior to receiving his M.D. from the Indiana University School of Medicine, Indianapolis, IN. He then went on to complete an internal medicine residency at the University of Colorado School of Medicine, Aurora, CO, as well as infectious disease and medical microbiology fellowships at Emory University School of Medicine, Atlanta, GA. Since July 2016, Dr. Lutgring has been an assistant professor of medicine at Emory University.



Jonathan C. Gullett, M.D., completed his medical education at the University of Central Florida College of Medicine. He completed his residency at the Medical University of South Carolina and Fellowship at Emory University. Dr. Gullett has been in the field for over 4 years and is currently the Physician Director of Microbiology at Kaiser Permanente (Southern California Permanente Medical Group) Regional Reference Laboratories. He is interested in *Clostridium difficile* because it is the most common microbial cause of nosocomial infections in the United States and costs many billions each year in excess health care costs. Despite the advent of rapid, highly sensitive molecular tests, diagnostic challenges persist due to noninconsequential rates of colonization and the seemingly never-ending question of where to draw the proverbial line.



Cassiana E. Bittencourt received a doctor of medicine degree in Brazil (2003). She completed her clinical pathology residency in 2008 at the University of Sao Paulo. After receiving her education in Brazil, she moved to the United States, where she finished residency training in pathology at the University of South Florida (2015) and microbiology fellowship at UT Southwestern (2016). She is certified by the American Board of Pathology in Anatomic and Clinical Pathology as well as Medical Microbiology. In August 2016, she joined the Department of Pathology at UC Irvine as Medical Director of Microbiology and serves as an Assistant Professor of Clinical Pathology in the UC Irvine School of Medicine. One of her clinical interests is the application of non-culture-based diagnostic methods in microbiology.



Susan Benson graduated in medicine from the University of Queensland before completing specialty training to become a Fellow of the Australasian Colleges of Physicians (Infectious Diseases) and Pathologists (Microbiology) and the College of Health Service Management. Over the last 20 years, she has worked as an infectious disease physician and in diagnostic medical microbiology, combined with teaching and research with the University of Western Australia. In 2014, she was awarded a Churchill Fellowship to examine the role of microbiology service delivery, education, and informatics to support better antimicrobial prescribing. Currently, Dr. Benson is employed by the Western Australian Department of Health, PathWest Laboratory Medicine, Perth, as clinical lead in a complex system improvement initiative to improve the management of patients with infection. She is also a Clinical Associate Professor at the University of Western Australia.



April M. Bobenchik received her undergraduate degree in medical technology from Marist College and graduate degrees (M.S. and Ph.D.) in biomedical science from Northeastern University and the University of Connecticut Health Center, respectively. She completed a CPEP-accredited medical microbiology and public health fellowship at the David Geffen Medical School, UCLA, Los Angeles, CA, and is a diplomate of the American Board of Medical Microbiology. Currently, Dr. Bobenchik is the Associate Director of Microbiology and Infectious Diseases Molecular Diagnostics at Lifespan Academic Medical Center and Assistant Professor of Pathology and Laboratory Medicine at Warren Alpert Medical School of Brown University, Providence, RI. She has a specific interest in education and advocates for the use of evidence-based best practices to help guide laboratorians in their daily practices and to help educate clinicians in appropriate test selection. Dr. Bobenchik has been a medical microbiologist for the past 4 years.



Robert L. Sautter obtained his B.S. and M.S. at Eastern Michigan University and completed his Ph.D. in microbiology from Wayne State University in 1982. He is the principal for RL Sautter Consulting LLC in South Carolina. Past affiliations include Director of Microbiology and Point of Care for Carolinas Pathology Group, Director for a County Health Laboratory, Director of the School of Medical Technology for Harrisburg Hospital, and Assistant Professor for Penn State Hershey Medical Center. Dr. Sautter also served as a CLIA medical laboratory director for 20 years, from 1995 to 2015. Dr. Sautter served on the North Carolina Response Advisory Group and HAI committee. He is also a member of the Institutional Biosafety Committee (IBC) for the University of North Carolina—Charlotte. Dr. Sautter served as clinical microbiology Chair of the American Society for Microbiology. He also served on the Clinical Laboratory Improvement Advisory Committee and Board of Scientific Counselors OIA at the CDC.



Vickie Baselski received her doctoral degree in microbiology from the University of Texas at Austin in 1978 and completed a postdoctoral fellowship in medical and public health laboratory microbiology at the Centers for Disease Control. She is a diplomate of the American Board of Medical Microbiology and a fellow in the American Academy of Microbiology. Her professional service activities have included technical direction of microbiology at the University of Tennessee Bowld Hospital, the Regional Medical Center at Memphis, and the Memphis Pathology Laboratory/American Esoteric Laboratory. She currently provides technical direction for microbiology at Methodist University Hospital and Regional One Health and is laboratory director at the Shelby County Health Department. She is a professor of pathology and laboratory medicine at the University of Tennessee Health Science Center. She served as a member and then Chair of the Professional Affairs Committee, where she has presented to federal, state, and private professional groups in areas relating to reimbursement for microbiological testing, and is currently a member of the CMS Advisory Panel on Clinical Diagnostic Laboratory Tests.



Michel C. Atlas was the librarian on this project. She obtained her undergraduate education at the State University of New York at Stony Brook and earned a master's degree in library service from Rutgers, the State University of New Jersey. For over 26 years, she served as a Reference Librarian at Kornhauser Health Sciences at the University of Louisville.



Elizabeth M. Marlowe received her Ph.D. from the University of Arizona in 1999 and completed a Clinical Fellowship in 2001 at the UCLA David Geffen School of Medicine, after which she continued at UCLA as an Adjunct Professor in the College of Medicine and as a Research Coordinator at the Wadsworth Anaerobe Laboratory at the West Los Angeles VA. In 2002, she joined Gen-Probe Incorporated, working on rapid molecular diagnostics. From 2005 to 2016, she provided laboratory leadership within the Kaiser Permanente Regional Laboratories in both the Southern California region, as the Assistant Director, and the Northern California region, as Technical Director over Microbiology. She is currently Director of Medical Affairs for Microbiology in the global division of Roche Molecular Systems. With over 18 years of experience, her research areas of interest are focused on the translation of new infectious disease diagnostics for use in evidence-based research for clinical microbiology practice.



Nancy S. Miller, M.D., F.C.A.P., F.A.S.C.P., received her M.D. from the State University of New York at Stony Brook School of Medicine. She completed an internship and residency in Anatomic and Clinical Pathology and a fellowship in Medical Microbiology, all at the Johns Hopkins Medical Institutions. Dr. Miller has been a practicing clinical microbiologist since 2001. She is currently the Medical Director of Clinical Microbiology and Molecular Diagnostics at Boston Medical Center and Associate Professor, Department of Pathology and Laboratory Medicine, at Boston University School of Medicine in Boston, MA. Prior to this, Dr. Miller was an Assistant Professor, Department of Pathology and Laboratory Medicine, at Boston University School of Medicine as well as Attending Pathologist in Anatomic and Clinical Pathology and Medical Director in the Microbiology Laboratory at Washington Hospital Center in Washington, DC. Dr. Miller is interested in the *C. difficile* guidelines because she wants to support efforts that improve patient care by fundamentally redefining what we do and how we do it via well-designed laboratory studies and meaningful evidence-based analysis.



Continued next page

Monika Fischer received her medical degree from Semmelweis University, Budapest, Hungary, in 1992. After completing her training in Internal Medicine and Gastroenterology, she moved to the United States and worked as a postdoctoral fellow at Indiana University. She completed fellowships in gastroenterology and inflammatory bowel diseases (IBD) at Indiana University in 2011, earned a master's of science in Clinical Science, and joined Indiana University Health as faculty. Currently, she is an Associate Professor of Medicine and Director of the IBD Program. As a health care provider to many IBD patients, she recognized the magnitude of *C. difficile*-related harm early in her career. She has built and is currently directing the Fecal Microbiota Transplant Program, through which over 700 patients were successfully treated. Her research focuses on outcomes of fecal microbiota transplantation in *C. difficile* infection, with emphasis on severe, therapy-refractory cases.



Sandra S. Richter received her M.D. from the University of Missouri—Columbia in 1996. She completed clinical pathology residency and microbiology fellowship training at the University of Iowa. She served as Director of the Microbiology Laboratory at the University of Iowa Hospital and Clinics before moving to the Cleveland Clinic, where she is the Section Head of Microbiology and an Associate Professor of Pathology at the Cleveland Clinic Lerner College of Medicine of Case Western Reserve University. Dr. Richter's research interests include the evaluation of new diagnostic tests, outcome-based laboratory utilization, and the epidemiology of antimicrobial resistance.



Peter Gilligan, Ph.D., D.(A.B.M.M.), F.(A.A.M.), is the Director, Clinical Microbiology-Immunology Laboratories, UNC HealthCare, and Professor of Microbiology-Immunology and Pathology-Laboratory Medicine, UNC School of Medicine. He received his Ph.D. in Microbiology from the University of Kansas. He completed his postdoctoral fellowship in medical and public health microbiology in 1980 and published his first paper on *Clostridium difficile* in 1981. Dr. Gilligan is the coauthor of the 2013 American College of Gastroenterology's *Guidelines for Diagnosis, Treatment and Prevention of Clostridium difficile Infections*.



James W. Snyder is a professor of pathology and laboratory medicine in the Department of Pathology and Laboratory Medicine, University of Louisville, where he serves as the director of clinical microbiology and molecular diagnostics at the University of Louisville Hospital, Louisville, KY. He received his Ph.D. in Biology (Clinical Microbiology) from the University of Dayton. He is a Diplomate of the American Board of Medical Microbiology and is a Fellow in the American Academy of Microbiology. His former positions include the Chief of Microbiology at the Veterans Administration Hospital, and Norton-Children's Hospital, Louisville, KY. His activities as a member of the American Society for Microbiology include being former chairperson of Division C and serving on the respective Professional Affairs and Laboratory Practices Committees. He served as ASM's representative to the Association of Public Health Laboratories (APHL) and coordinating editor for the Laboratory Response Network (LRN) Sentinel Level Protocols for Biothreat Agents and is a former member of the APHL Biosafety and Biosecurity Committee in addition to having been a member of the Sentinel Laboratory Partners Outreach Committee. He currently serves on the editorial boards of the *Journal of Clinical Microbiology* and the *Journal of Surgical Infections*. He was elected as President-Elect of the South Central Association for Clinical Microbiology (SCACM) and will serve as President for the year 2018.

